

**Scale, scaling and multifractals  
in geophysics**

**Part 2:  
Fractal sets, multifractal cascades**

6 May, 2014

Course at U. Paris Sud, May 6, 7 2014



# Scale Invariance sets and fields

Course at U. Paris Sud, May 6, 7 2014

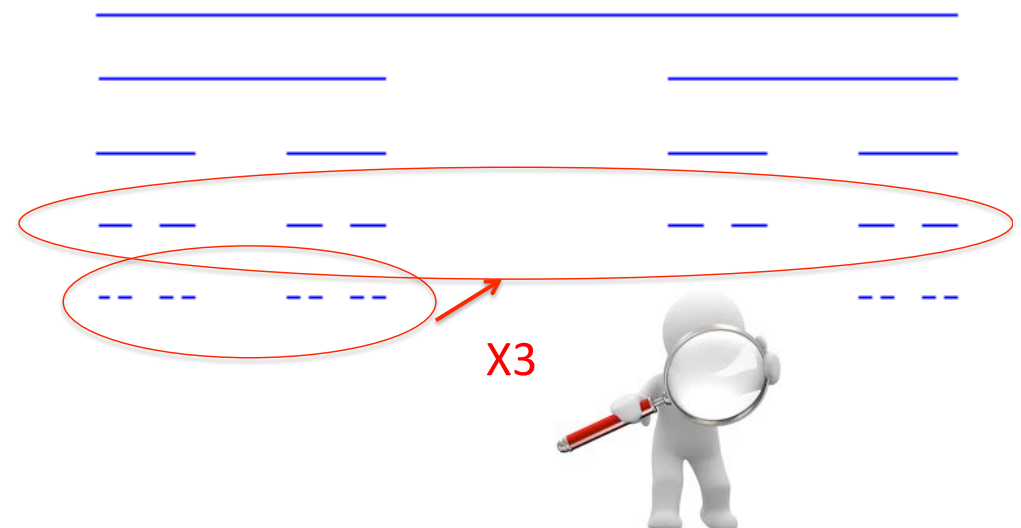
# Scale invariant geometric sets: Fractals

The simplest fractal, the Cantor set (1871)

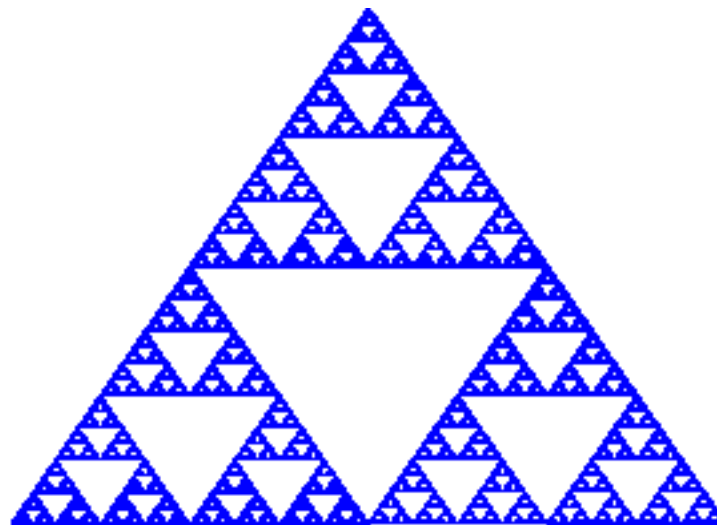
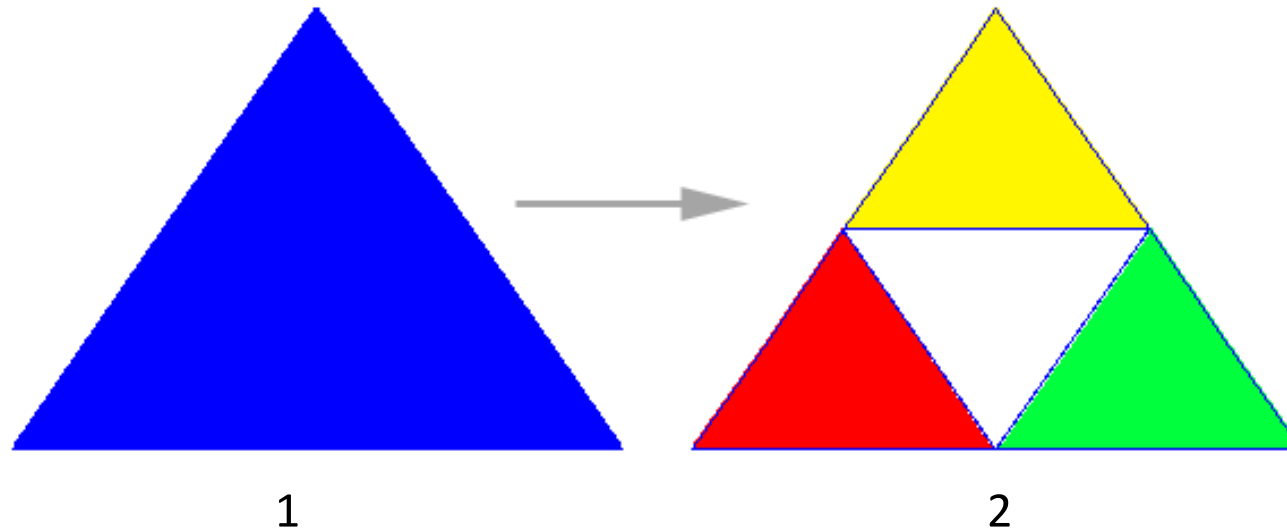
- Start with:



iterate:



# Sierpinski Triangle

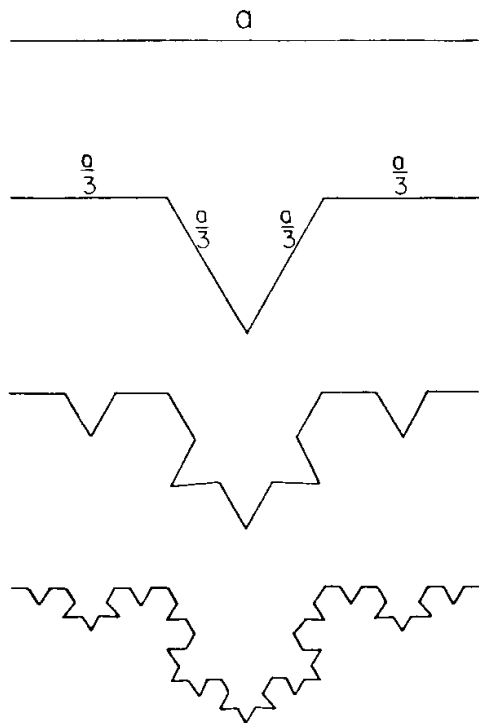


**10 iterations**

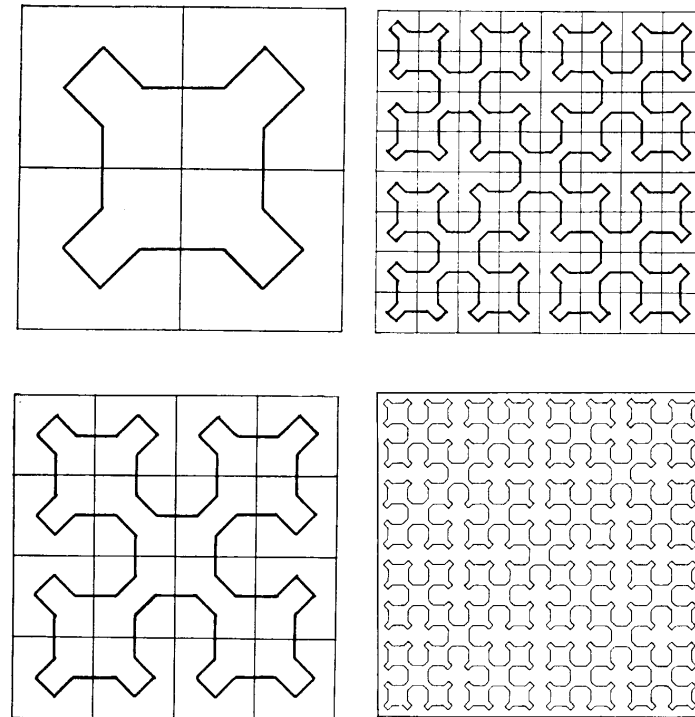


# Early Geophysical applications of Fractal sets

Set: Black / white, single fractal dimension



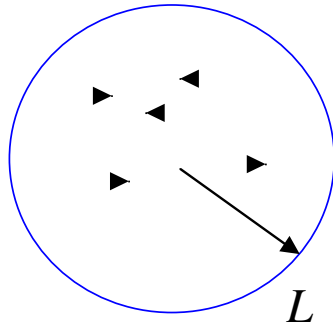
A fractal Koch curve ([*Koch*, 1904]), reproduced from [*Welander*, 1955] to illustrate the mixing of a two dimensional fluid.



A fractal Peano curve, reproduced from [*Steinhaus*, 1960] showing how a line (dimension 1) can literally fill the plane (dimension 2), illustrating how streams can fill a surface.

# Isotropic Scale Invariance and fractal sets

Fractal Dimension:



$$n(L) \propto L^D$$

Number of points

$$\rho(L) = \frac{n(L)}{L^d} \propto L^{D-d} = L^{-C}$$

Density of points

d=dimension of space  
D= fractal dimension of set  
C=d-D= fractal codimension

Scale invariance:

$$n(\lambda L) = \lambda^D n(L)$$

**D=scale invariant**

Same form after zoom by factor  $\lambda$ .

# Meteorological measuring network

Fractal set: each point is a station

9962 stations (WMO, 1986)



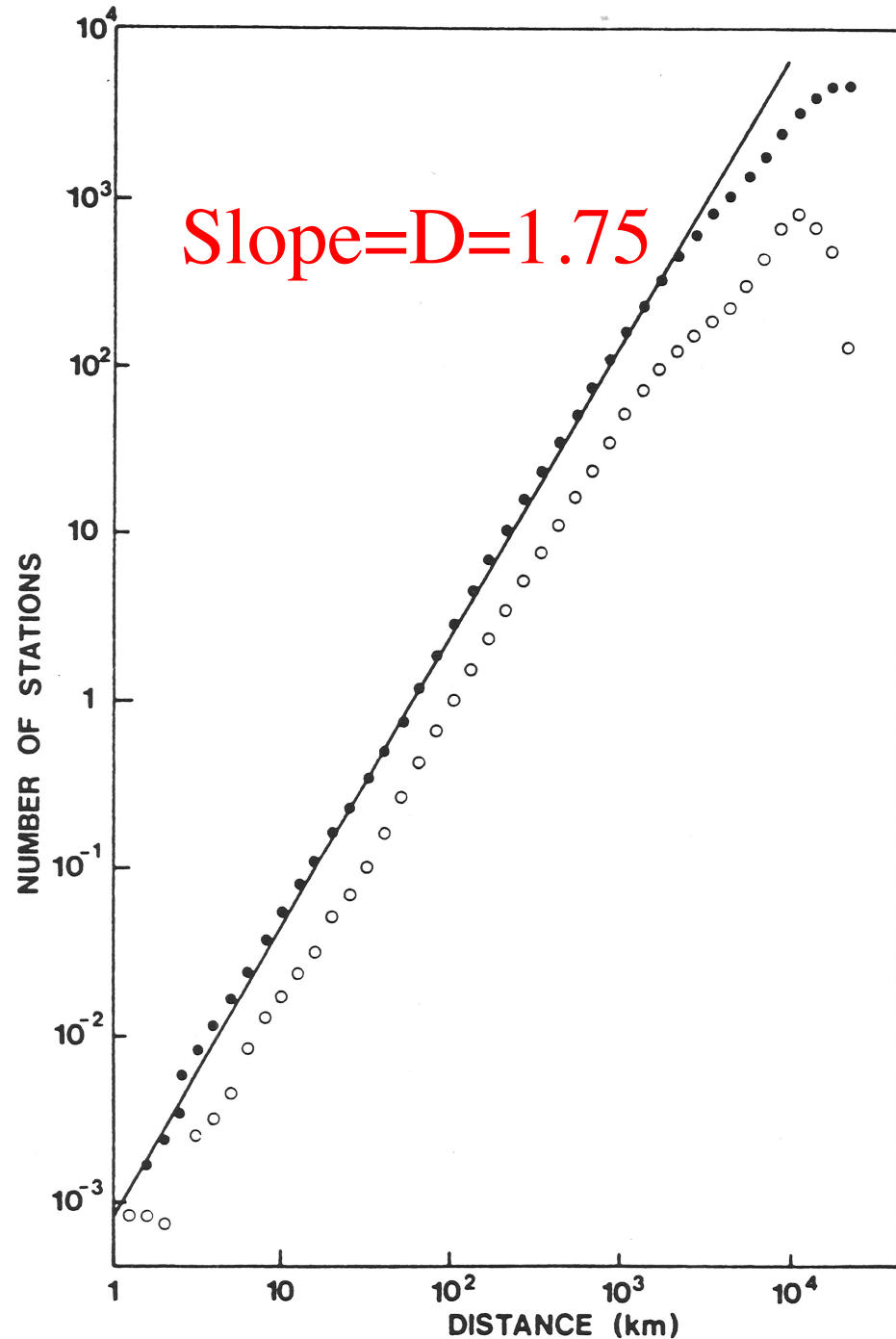
Number

$$n(L) \propto L^D$$

Density

$$\rho(L) = n(L)L^{-2} \propto L^{-C}; \quad C = d - D; \quad d = 2$$

The fractal dimension of the network=  
1.75





# (1) Fractal Codimensions: Geometric

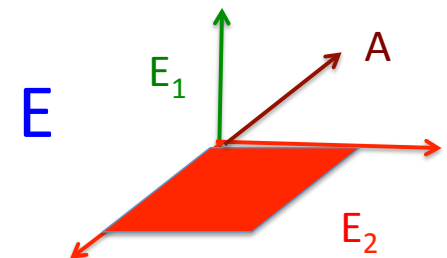
The notion of fractal codimension  $C_F$  can be defined both statistically and geometrically. The former is more useful and general since it applies not only to deterministic but also to stochastic processes.

Let  $A \subset E$  (the embedding space) with  $\dim(E) = D$  and  $\dim(A) = D_F(A)$ . Then the codimension  $C_F(A)$  is defined as:

$$C_F(A) = D - D_F(A)$$

This definition corresponds merely to an extension of the (integer) codimension definition for vector sub-spaces, i.e.,  $E_1$  and  $E_2$  being in direct sum (i.e.,  $E_1 \cap E_2 = \emptyset$ ):

$$E = E_1 \oplus E_2 \Rightarrow \text{codim}(E_1) = \dim(E_2)$$



Example:  $E_1$  = line,  $E_2$  = plane,  $E$  = 3-D space

a line ( $D_F(A)=1$ ) in three dimensional space:  $\dim(E)=D=3$ , hence:  $C_F(A)=3-1=2$

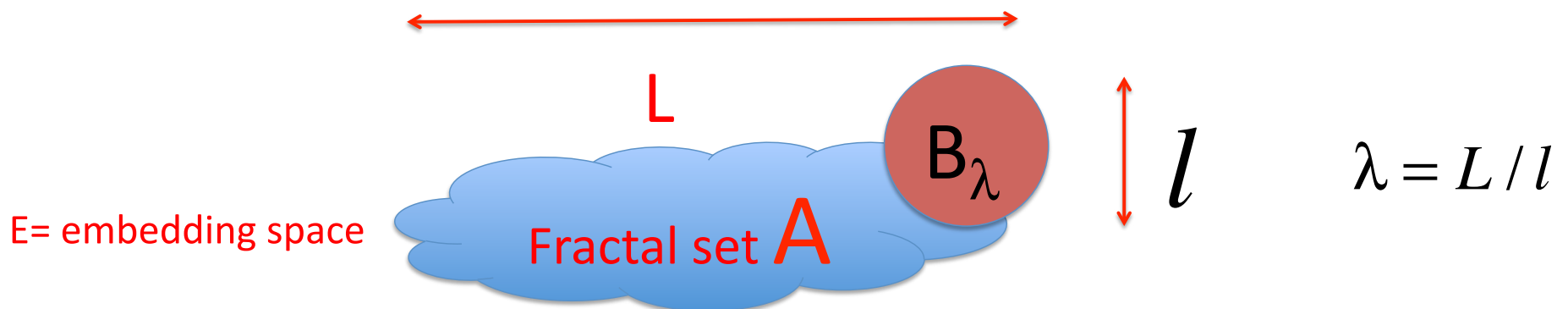
## (2) Fractal Codimensions: Probabilistic

The codimension  $C_F$  can be introduced directly.

Consider the (scaling) behaviour the probability (“Pr”) that a ball  $B_\lambda$  (of size  $\ell = L / \lambda$ ) intersects the set  $A$  is:

$$\Pr(B_\lambda \cap A) \sim \lambda^{-C_F(A)}$$

where  $B_\lambda$  = ball of size and  $\ell = L / \lambda$  and  $C_F$  is thus directly defined as an exponent measure of the fraction of the space occupied by the fractal set  $A$  (size  $L$ ) in an embedding space  $E$  which can even be an infinite dimensional space.



# Geometric versus probabilistic

## Relating the two definitions

Since the probability of the event  $(B_\lambda \cap A)$  is defined as:

$$\Pr(B_\lambda \cap A) \sim \frac{N(B_\lambda \cap A)}{N(B_\lambda \cap E)} \sim \frac{\lambda^{D_F(A)}}{\lambda^{D(E)}}$$

Number of balls  $B_\lambda$  needed to cover  $A$

Number of balls  $B_\lambda$  needed to cover  $E$

where  $N(B_\lambda \cap A)$  refers to for example the number of balls  $B_\lambda$  needed to cover the set  $A$  and  $N(B_\lambda \cap E)$  is the corresponding number for the entire space. It is easy to check that when  $C_F(A) < D = \dim(E) < \infty$  the two definitions are equivalent:

$$C_F(A) \leq D < \infty, \{ \text{definition 1} \equiv \text{definition 2} \} \quad \forall D_F \geq 0$$

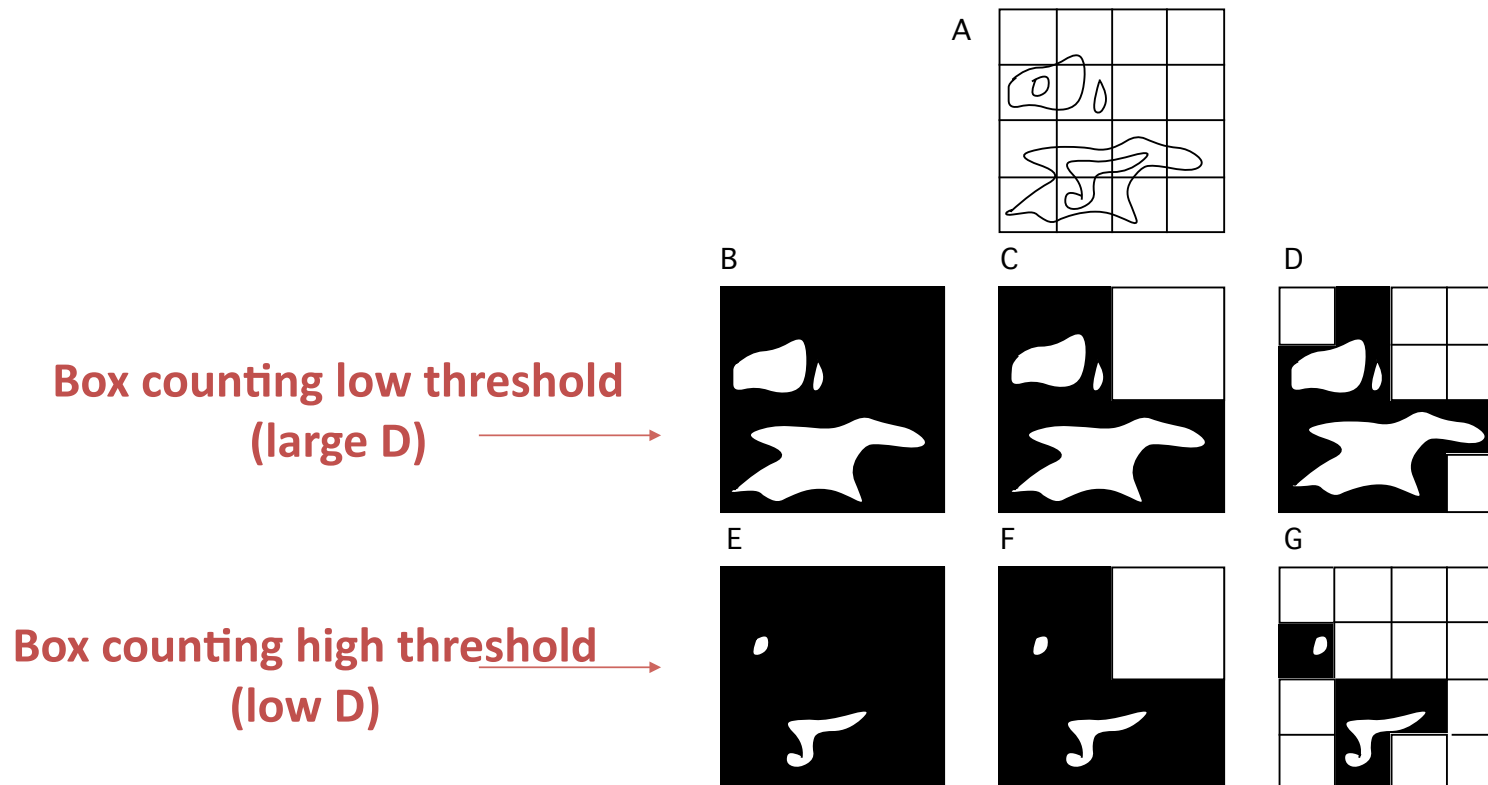
However, when  $C_F(A) > D$ , then they no longer agree since it implies  $D_F(A) < 0$  which is impossible.

# Multifractal fields: Cascades and Multifractals



# Multifractality and Functional Box Counting

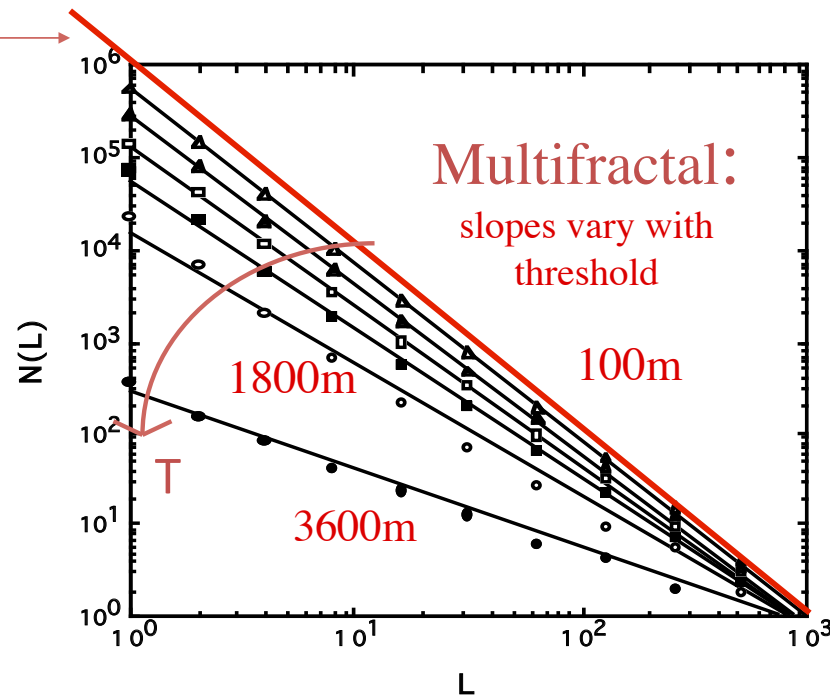
$$N_T(L) \approx L^{-D(T)}$$



- Monofractal:  $D(T) < 2$ , constant
- Multifractal:  $D(T) < 2$ , decreasing

# Functional box counting on French topography: 1 -1000km

Slope =2  
(required for classical geostatistics - regularity of Lebesgue measures)



$$N_T(L) \approx L^{-D(T)}$$

Systematic resolution dependence

km

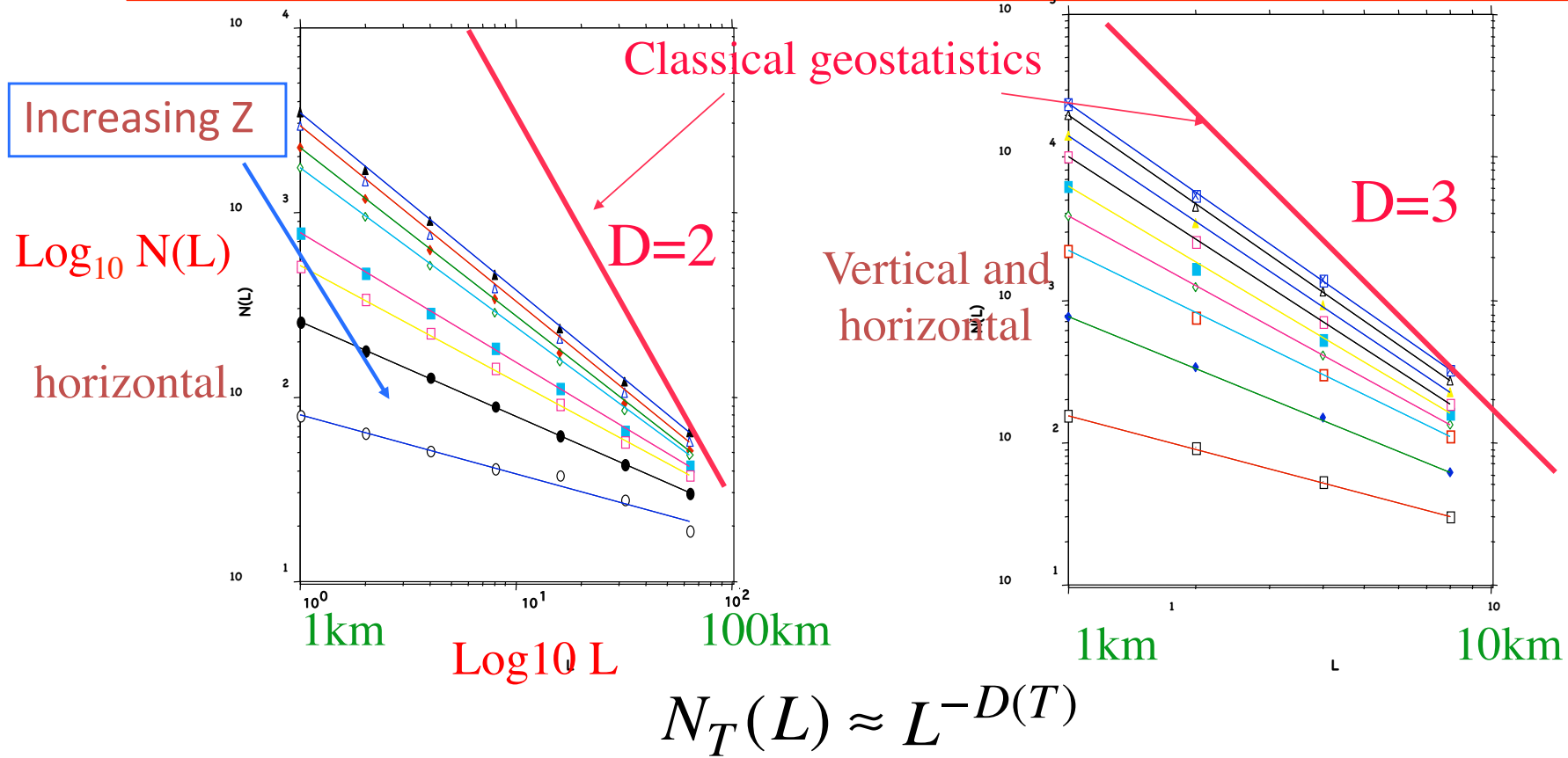
## Implications for geostatistics:

The areas  $A_T$  exceeding a given threshold decrease as the resolution becomes finer (decreasing  $L$ ):

$$A_T = L^d N_T = L^{C(T)}; \quad C(T) = d - D(T)$$

Unless  $C(T) = 0$ , the areas depend on the subjective resolution  $L$ ; the reference lines indicate that for the topography, all the regions defined by the thresholds have  $C(T) = d - D(T) > 0$  so that they have systematic resolution dependencies.

# Functional Box counting on 3D radar rain scans



Radar reflectivity thresholds increasing (top to bottom) by factors of 2.5 (data from Montreal).

# Aircraft temperature transect (12km altitude)

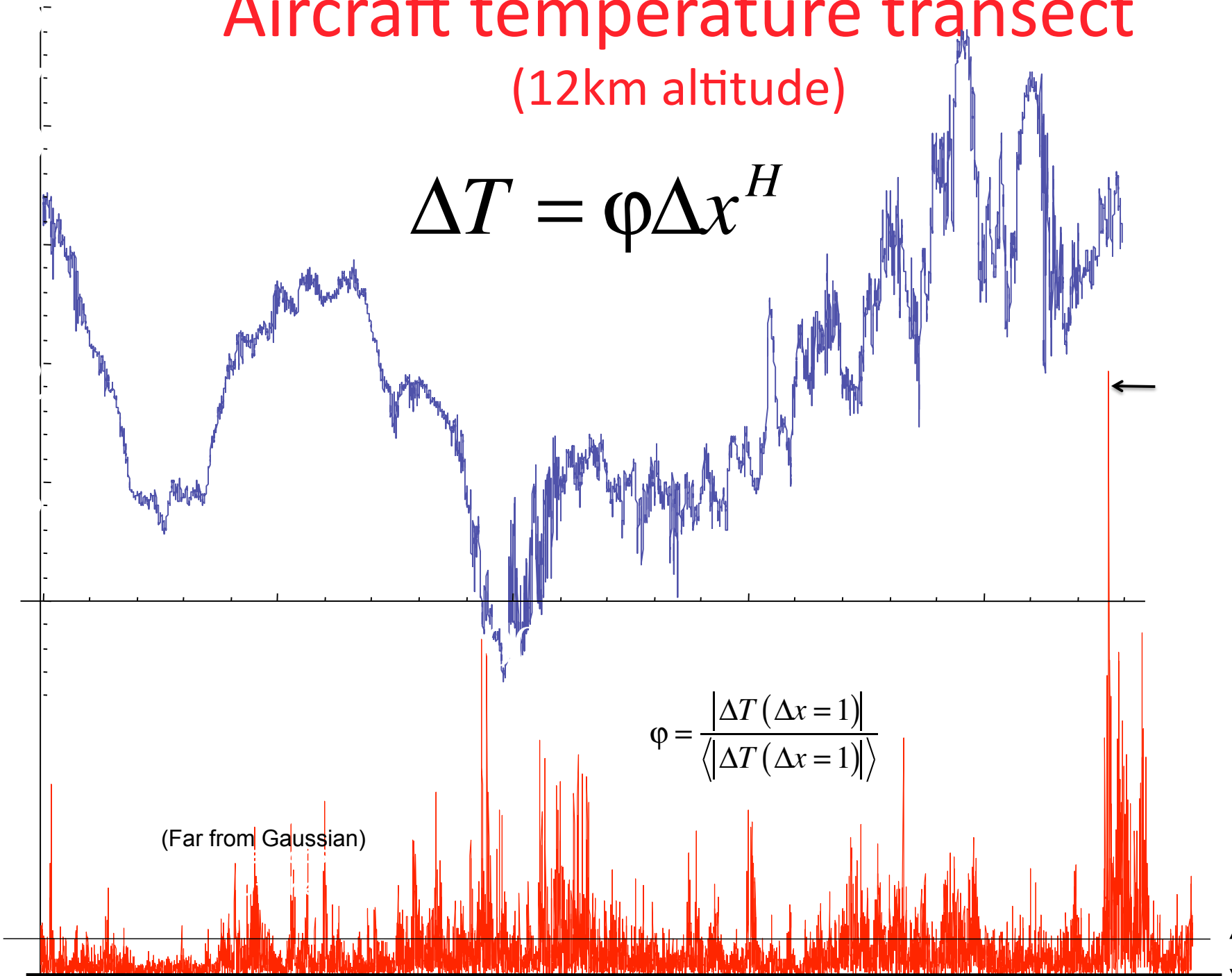
$$\Delta T = \varphi \Delta x^H$$

$\varphi$

(Far from Gaussian)

$$\varphi = \frac{|\Delta T(\Delta x = 1)|}{\langle |\Delta T(\Delta x = 1)| \rangle}$$

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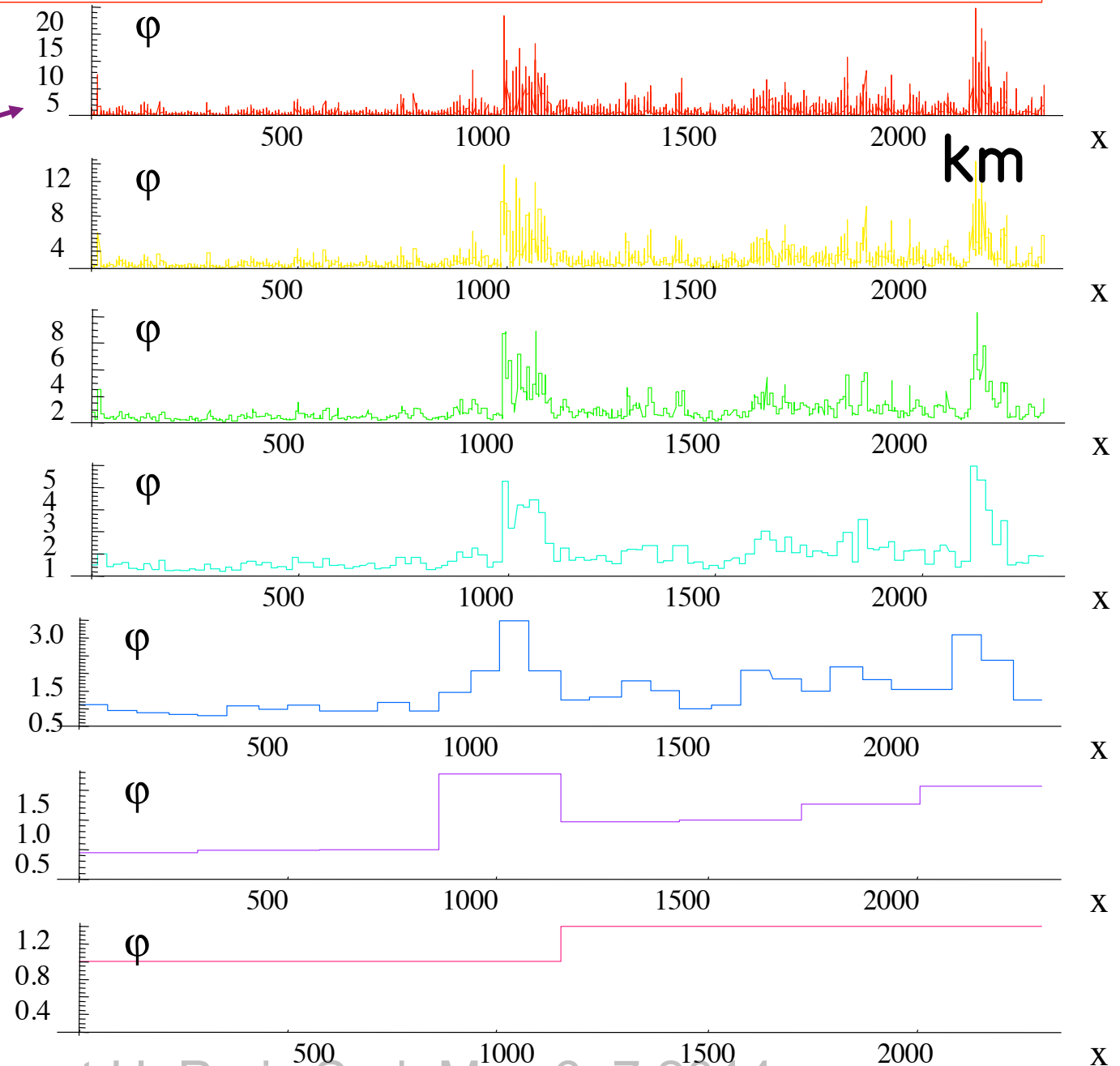




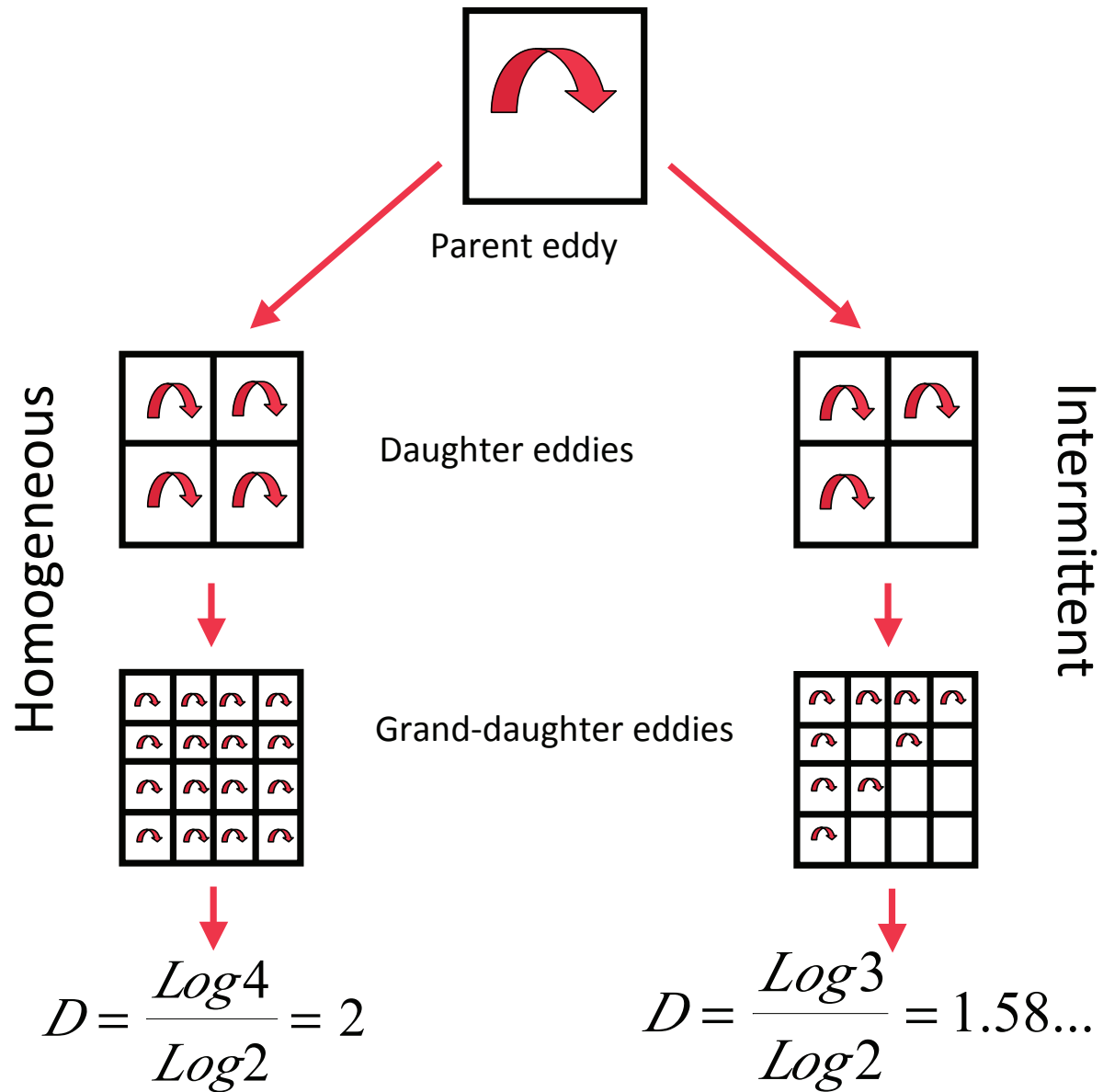
# Degrading the resolution

Temperature  
turbulent flux  $\phi$   
at 280m resolution

High to low  
Resolution:  
degrading by  
factors of 4



# Cascades



# Beta model

An initial attempt to handle intermittency reduces it to the simple notion of “on/off” intermittency, i.e. a cascade with the simple alternative alive/dead of the offspring.

This leads to a confinement of the turbulence to a tiny support; a very small subregion of the flow. The right hand side of the figure shows the result of such a stochastic cascade obtained by randomly multiplying the energy flux of a “mother” eddy to obtain that of the “daughter” eddies either by 0 (dead sub-eddy) or by a positive value  $\lambda_0^c$

(corresponding to an active sub-eddy, with fixed probability  $\lambda_0^{-c}$

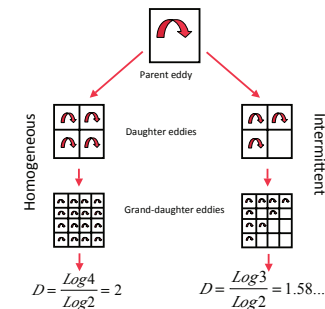
In this model, we divide the spatial scales by  $\lambda_0$  (here  $\lambda_0 = 2$ ) and then flip coins to determine the on or off state; more precisely:

Each step:

$$\Pr(\mu\varepsilon = \lambda_0^c) = \lambda_0^{-c}$$

$$\Pr(\mu\varepsilon = 0) = 1 - \lambda_0^{-c}$$

After n steps: 
$$\varepsilon_n = \prod_{j=1}^n \mu\varepsilon_j$$

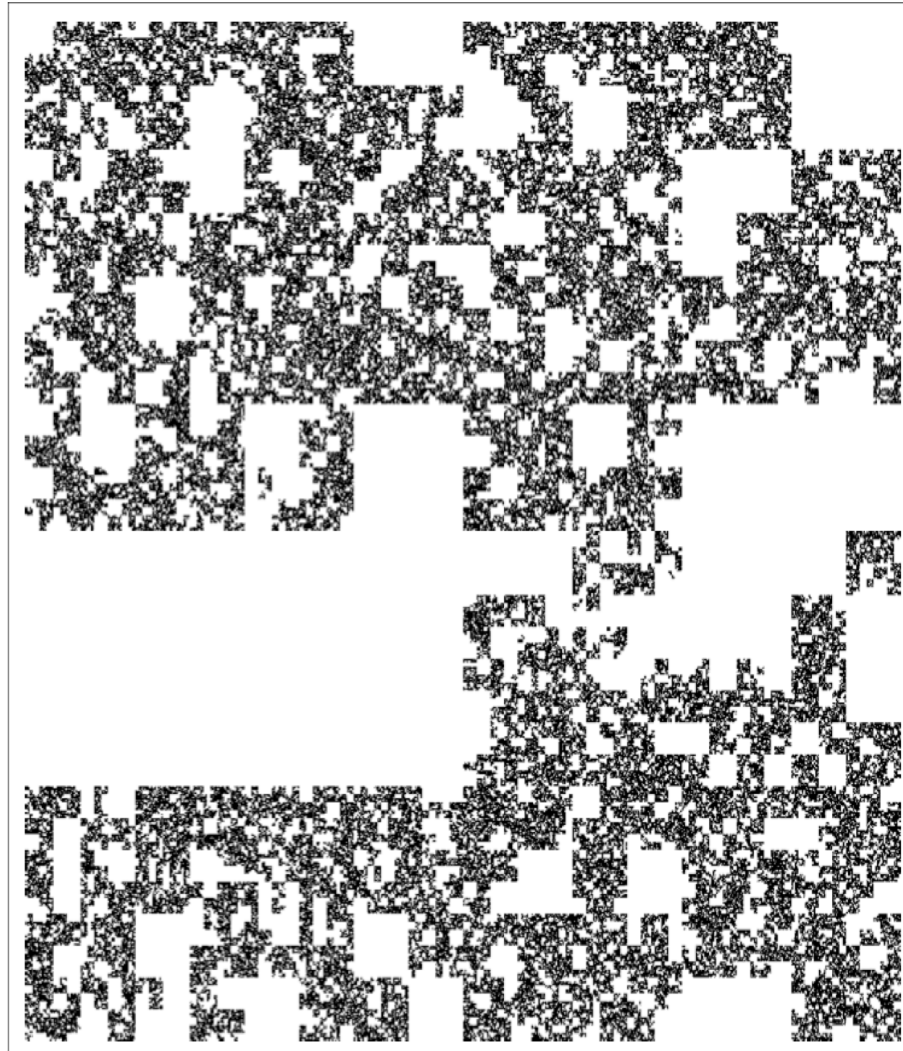


(“Pr” indicates “probability”). The nonzero value is taken as  $\mu\varepsilon = \lambda_0^c$  so that the mean  $\langle \mu\varepsilon \rangle = 1$  ; this implies a scale by scale conservation of the flux  $\varepsilon$ .

After n steps:  $\lambda = \lambda_0^n$      $\Pr(\text{alive}) = (\lambda_0^{-c})^n = \lambda^{-c}$     Relation to dimension:  $N_{\text{alive}} = N_{\text{tot}} \Pr = \lambda^d \lambda^{-c} = \lambda^D$ ;     $D = d - c$

# Beta model

In this example, the probability that an eddy will remain alive is  $\lambda_0^{-C} = 0.87$  (using the scale ratio at each step  $\lambda_0 = 4$  here and the codimension  $C = 0.2$ ).



# Alpha model

The  $\alpha$  model is a two state (binomial) process with  $\mu\varepsilon =$  either  $\lambda_0^{\gamma_+}$  or  $\lambda_0^{\gamma_-}$  where  $\gamma_+ > 0$  corresponds to a boost ( $\mu\varepsilon > 1$ ) and  $\gamma_-$  to a decrease ( $\mu\varepsilon < 1$ ).

$$\Pr(\mu\varepsilon = \lambda_0^{\gamma_+}) = \lambda_0^{-c}$$

$$\Pr(\mu\varepsilon = \lambda_0^{\gamma_-}) = 1 - \lambda_0^{-c}$$

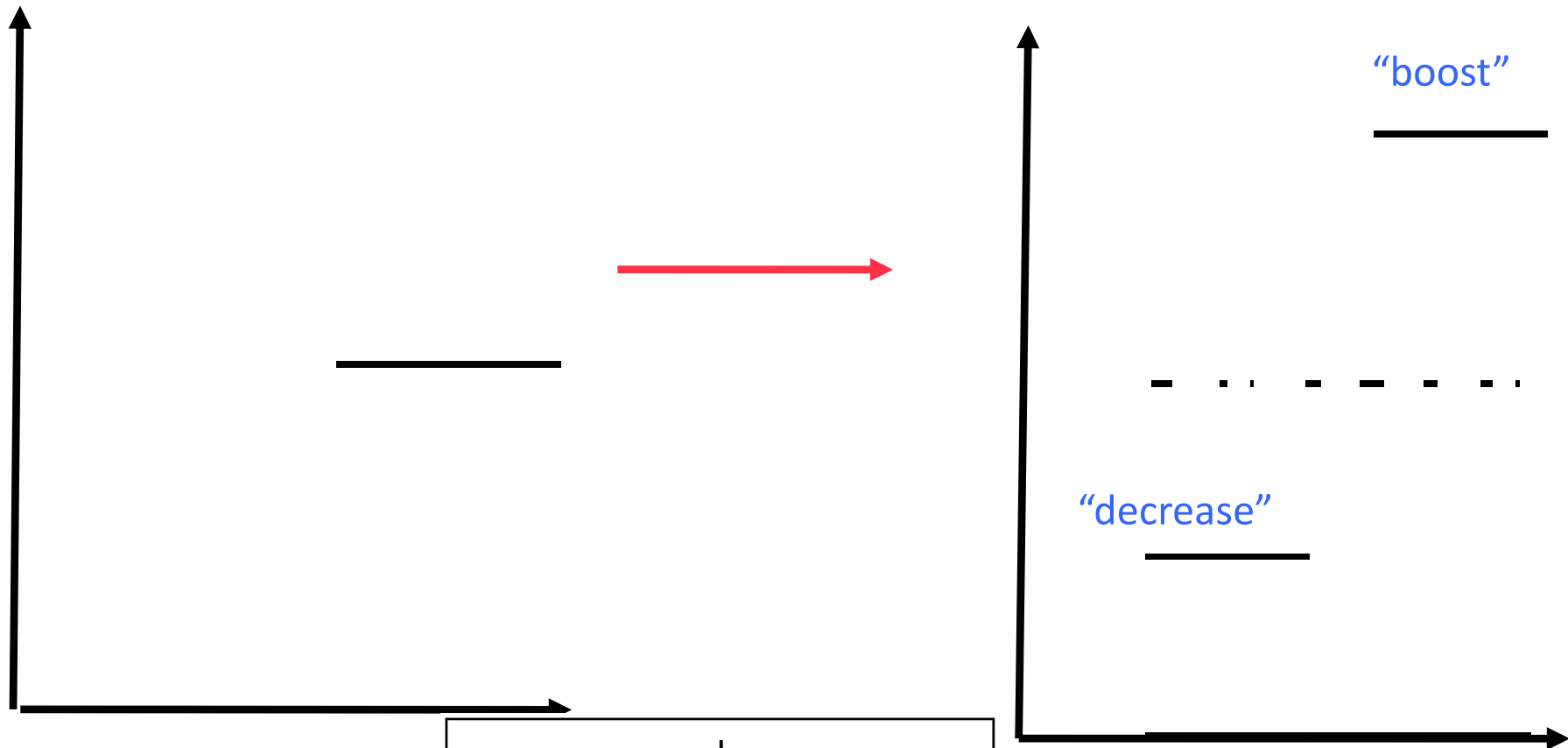
Although the  $\alpha$  model apparently involves three parameters ( $\gamma_+$ ,  $\gamma_-$ ,  $c$ ), due to the conservation constraint:

$$\langle \mu\varepsilon \rangle = \lambda_0^{-c} \lambda_0^{\gamma_+} + (1 - \lambda_0^{-c}) \lambda_0^{\gamma_-} = 1$$

We can see that the  $\beta$  model is recovered in the limit  $\gamma_- \rightarrow -\infty$  which is the same as  $\gamma_+ \rightarrow c$

# The $\alpha$ model

Simulations: **multiplicative** introduction of small scale details  
(low resolution to high)

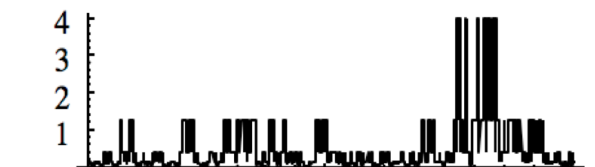
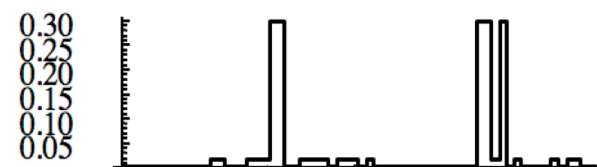
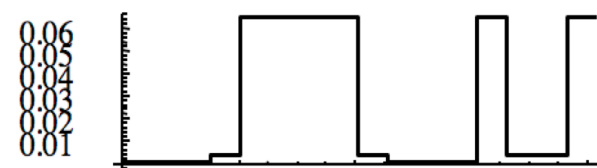
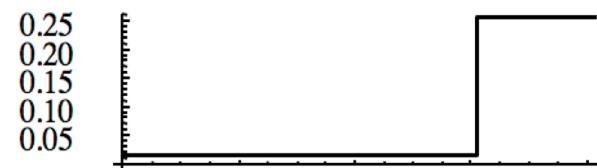
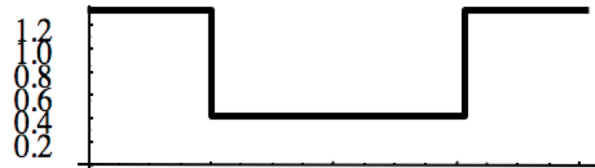


S+L 1983

# Alpha model

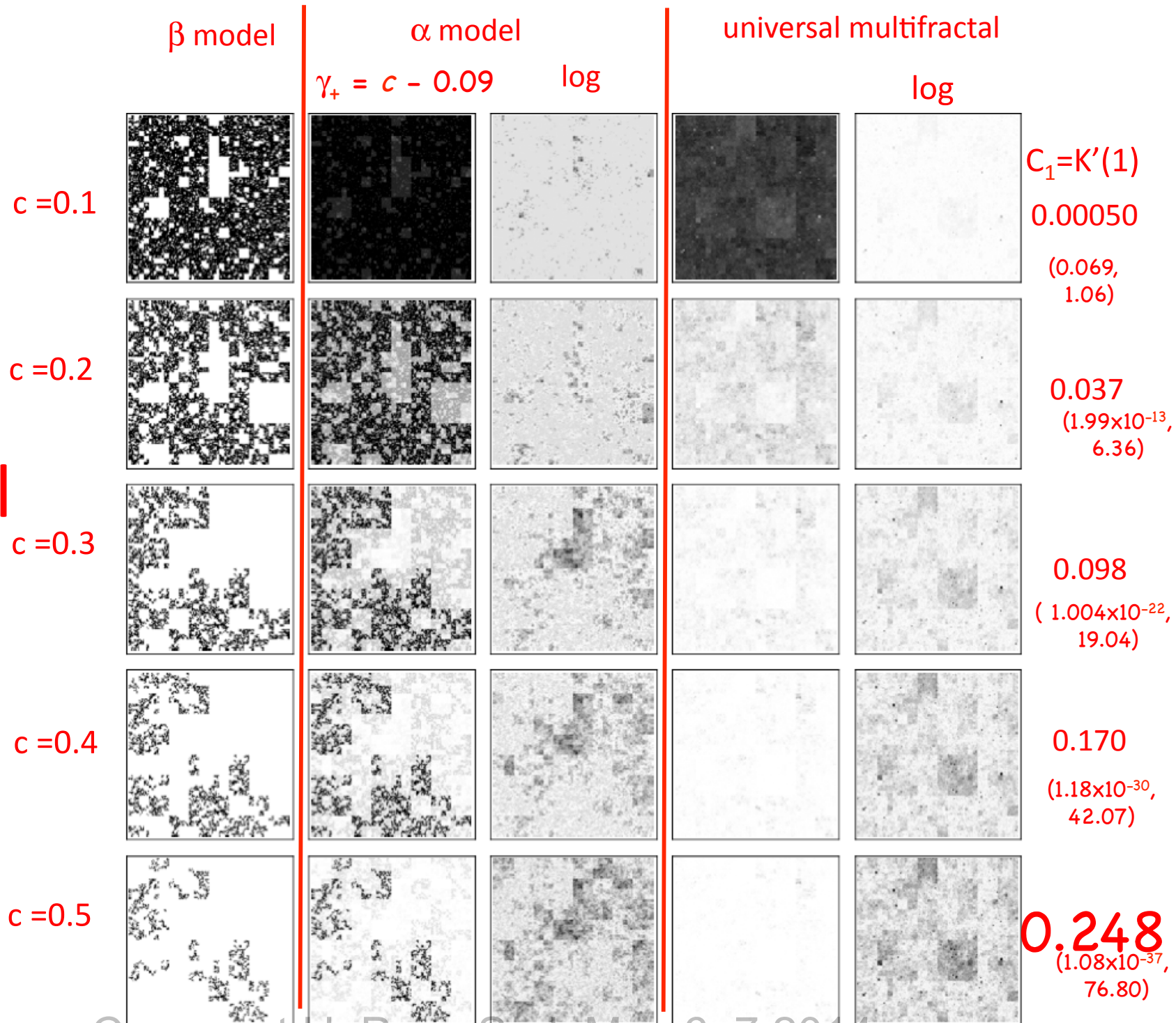
$\gamma_+ = 0.2, c = 0.3 (C_1 = 0.087)$

$\gamma_+ = 1.1, c = 1.2 (C_1 = 0.82)$



From top to bottom every second cascade step is shown (a factor of  $\lambda_0^2$  is shown, 10 steps in all, the total range of scales is  $2^{10} = 1024$ ). Notice the changing vertical scales

# 2-D Alpha model





# General cascade statistics

Characterize the statistics of  $\mu\varepsilon$  by  $K(q)$ :  $\langle \mu\varepsilon^q \rangle = \lambda_0^{K(q)}$  ← Scale ratio of each cascade step

← The notation “ $\mu$ ” indicating “multiplicative increment”; it is analogous to the use of the “ $\Delta$ ” to denote an additive increment.

$$\langle \varepsilon_n^q \rangle = \left\langle \prod_{j=1}^n \mu\varepsilon_j^q \right\rangle = \prod_{j=1}^n \langle \mu\varepsilon_j^q \rangle = \langle \mu\varepsilon^q \rangle^n = \lambda_0^{nK(q)}$$

We can now write the general expression for the statistical properties after a total scale range  $\lambda$ :

$$\langle \varepsilon_\lambda^q \rangle = \lambda^{K(q)}$$

Overall scale ratio since the cascade started:

$$\lambda = \lambda_0^n$$

This is the basic formula for cascade statistics. The specification of the statistics of  $\mu\varepsilon$ , and hence of  $\varepsilon_\lambda$  via statistical moments is equivalent to their specification by probabilities.

# The cascade generator: $\Gamma$

The overall characterization of the statistical properties is conveniently through the “moment scaling exponent”  $K(q)$ :

$$\begin{aligned} K(q) &= \text{Log}_{\lambda_0} \langle \mu \varepsilon^q \rangle = \text{Log} \langle \mu \varepsilon^q \rangle / \text{Log} \lambda_0 \\ &= \text{Log} \langle \mu \varepsilon^q \rangle^n / \text{Log} (\lambda_0)^n = \text{Log} \langle \varepsilon_\lambda^q \rangle / \text{Log} \lambda \end{aligned}$$

Introducing the (random) cascade “generator”  $\Gamma$ , the logarithm of the multiplier:

$$\Gamma = \text{Log} \varepsilon_\lambda$$

$K(q)$  is the (Laplace, base  $\lambda_0$ ) second characteristic function (“cumulant generating function”) of  $\Gamma$ :

$$K(q) = \log_\lambda \langle e^{q\Gamma} \rangle$$

# Examples of second characteristic Functions

Ex.1 Gaussian

$$\langle e^{qx} \rangle = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{qx} \underbrace{e^{-x^2/(2\sigma^2)}}_{p(x)} dx = e^{q^2\sigma^2/2}$$

$$K(q) = \log \langle e^{qx} \rangle = \frac{q^2\sigma^2}{2} \quad \text{Base e Laplace characteristic function}$$

Ex.2 Exponential:

$$p(x) = \frac{1}{2} e^{-|x|}$$

$$K(q) = -\log(1-q) - \log(1+q); \quad -1 \leq q \leq 1$$

# Properties of the Moment scaling exponent $K(q)$

- 1) In order to see the general shape of the  $K(q)$  function, we may first note that conservation from one scale to another requires  $K(1) = 0$ :

$$\langle \varepsilon_\lambda \rangle = 1 = \lambda^0 \quad \text{hence} \quad K(1) = \text{Log}_\lambda \langle \varepsilon_\lambda \rangle = \text{Log}_\lambda 1 = 0$$

- 2) In addition, because any positive number raised to the zero power is one, we have  $\langle 1 \rangle = 1$ , hence  $K(0) = 0$ .  $\langle \varepsilon_\lambda^0 \rangle = 1 = \lambda^0$

- 3) Finally, a basic property of second characteristic functions is that  $K(q)$  must be convex, i.e.  $K''(q) > 0$ ; this can be shown directly by doubly differentiating  $K(q) = \log \langle e^{q\Gamma} \rangle / \log \lambda$ .

The typical  $K(q)$  looks something like the next slide which shows the  $K(q)$  for the  $\alpha$  model and the universal multifractal models in the fourth and fifth columns of the earlier example. The models are tangent to each other at  $q = 1$  because the derivatives at  $q = 1$  were deliberately chosen to be equal to each other. This value:

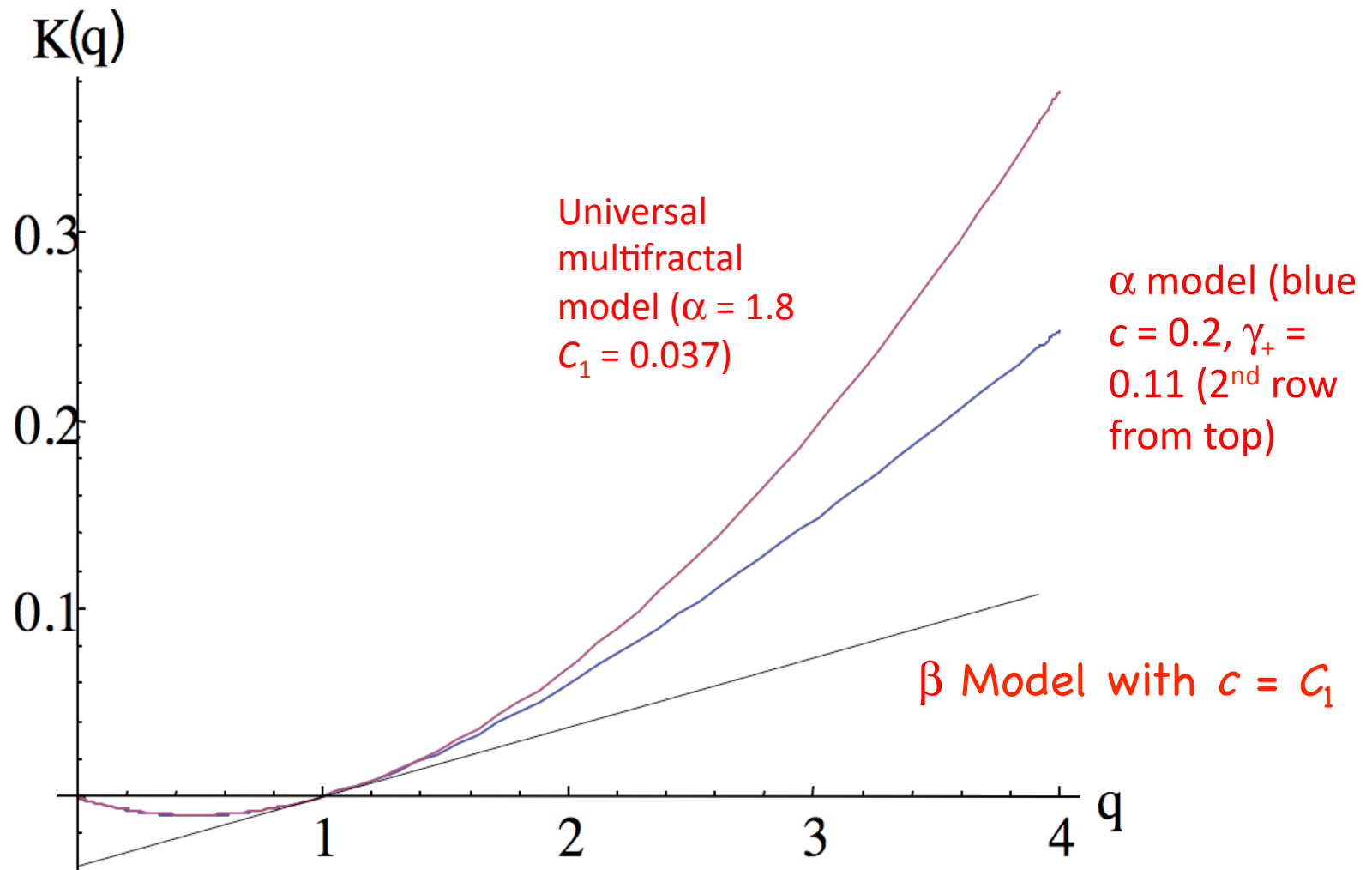
$$C_1 = K'(1) \quad \text{tangent at the mean}$$

$C_1 =$  “the codimension of the mean”; a characterization of the variability near the mean

We can already use this idea to give a “local” (in  $q$  space) definition of the “degree of multifractality”  $\alpha$ :

$$\alpha = K''(1) / K'(1) \quad \text{Curvature near the mean}$$

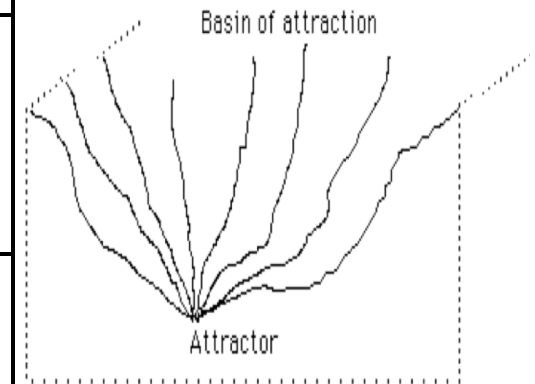
# Comparison of the $K(q)$ for examples



$$C_1 = K'(1)$$
$$\alpha = K''(1)/K'(1)$$

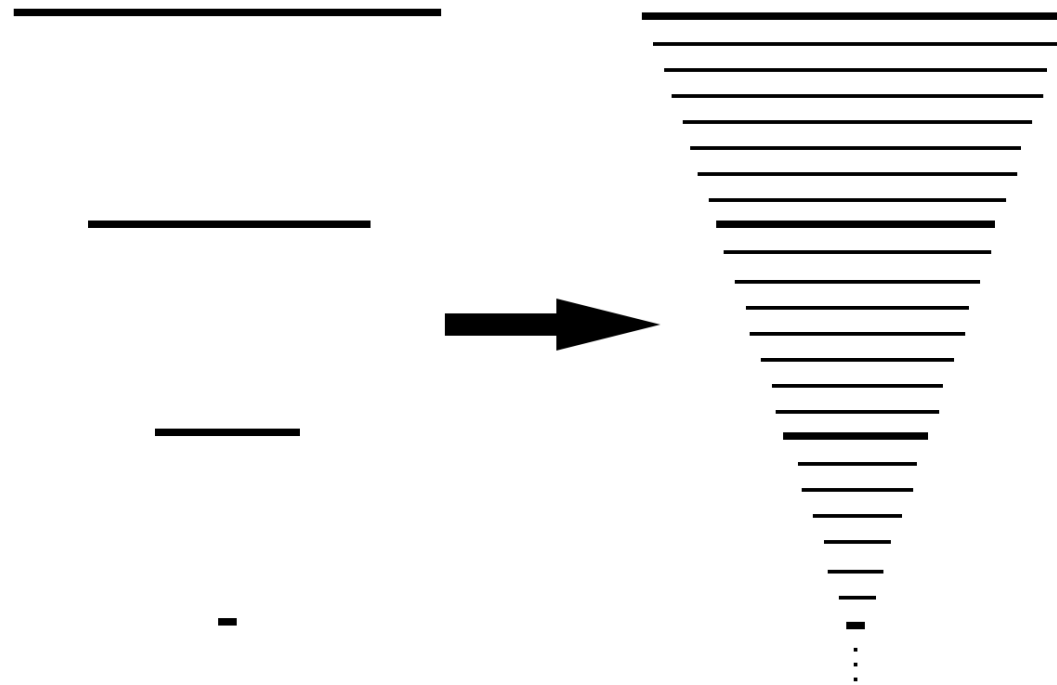
# Universality: How many parameters for turbulence?

<u>Answer</u>	<u>Date</u>	<u>References</u>	<u>Explanation</u>	<u>Parameters</u>
1	1941	Kolmogorov (Homogeneous turbulence)	$\Delta v_\lambda \approx \bar{\epsilon}^{1/3} \lambda^{-1/3}$	$H=1/3$
2	1962	Kolmogorov-Obukhov, (lognormal model)	$\langle \epsilon_\lambda^q \rangle = \lambda^{K(q)}$ $K(q) = \frac{\mu}{2}(q^2 - q)$	$H, \mu$
2	1964	Novikov-Stewart, Mandelbrot, Frisch et al, $\beta$ model	$K(q) = C_1(q - 1)$	$H, C_1$
$\infty$	1974	(Mandelbrot, 1974)	$K(q)$	Any $K(q)$ convex with $K(0)=K(1)=0$



# Routes to universality: 1) Densification of scales

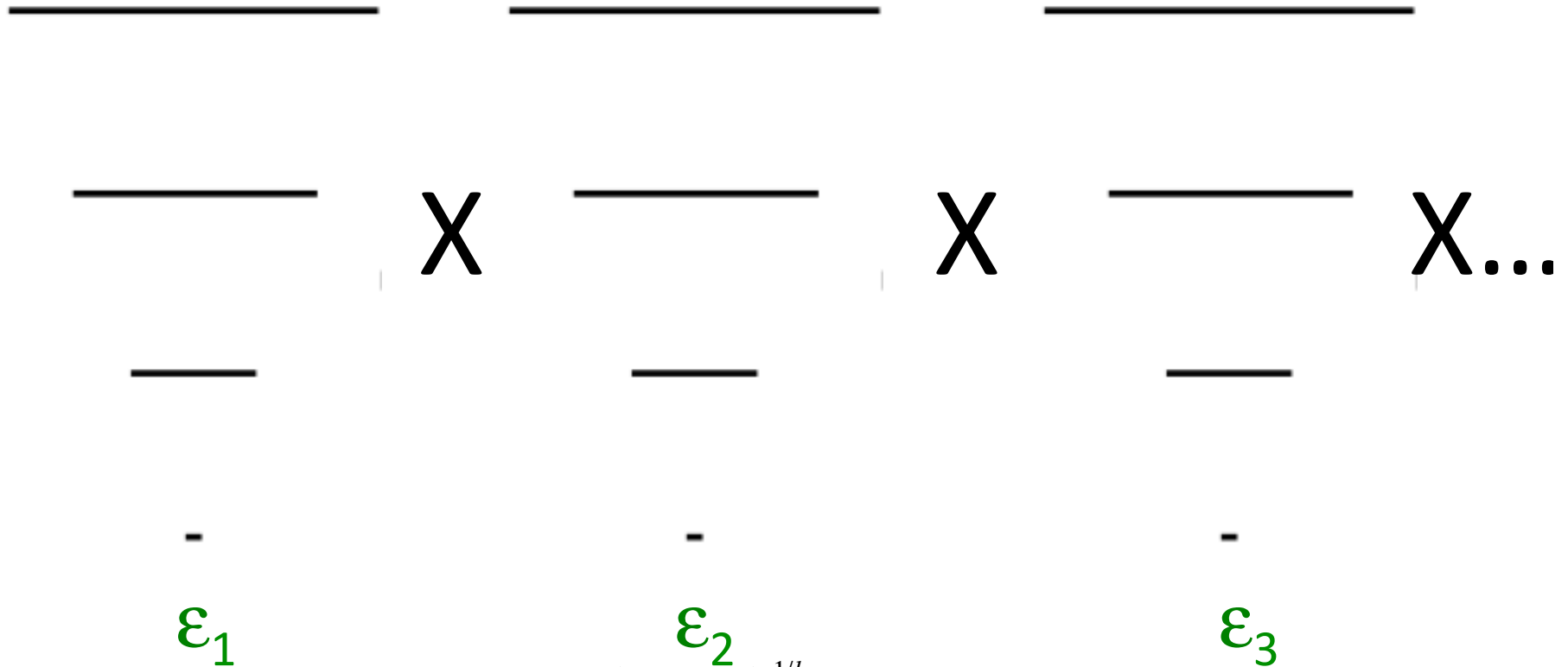
Discrete in scale  
(ex.  $\beta$ ,  $\alpha$  models)



Continuous in  
scale

# Routes to universality:

## 2) "Mixing" of independent discrete cascades



$$\varepsilon = \left( \prod_{i=1}^N \frac{\varepsilon_i}{a_N} \right)^{1/b_N}$$

N independent cascades,  
renormalized by  $a_N, b_N$

For the generators  $\Gamma = \log \varepsilon$

$$\Gamma = \frac{1}{b_N} \sum_{i=1}^N (\Gamma_i - \log a_N)$$

Normalized, centred sums



# Universality in cascades: a “multiplicative central limit theorem”

**Technical difficulty:** the cascade requires a scale by scale conservation principle, otherwise there are no well defined small scale cascade limits, and it turns out that this normalization is in contradiction with the normalization required for central limit convergence.

Cascade convergence:  $\langle \mu \varepsilon \rangle = 1$  hence  $\langle e^{\Delta \Gamma} \rangle = 1$

Central limit convergence:  $\langle \Delta \Gamma \rangle = 0$ , hence  $\langle \log \mu \varepsilon \rangle = 0$

Recall:  
 $\Delta \Gamma = \log \mu \varepsilon$

However, due to the convexity of the logarithm function, for any probability distribution of  $\mu \varepsilon$  which is constrained such that  $\langle \mu \varepsilon \rangle = 1$ , we have necessarily  $\langle \Delta \Gamma \rangle = \langle \log \mu \varepsilon \rangle < 0$

# Levy Generators (2)

The final normalization step needed for small scale convergence (analogous to the log-normal derivation:  $K(q) \rightarrow K(q) - qK(1)$ ) leads to:

$$K'(1) = A_\alpha (\alpha - 1) = C_1$$
$$K''(1) = A_\alpha \alpha (\alpha - 1) = \alpha K'(1) = \alpha C_1$$

The local (near the mean) curvature characterization is satisfied:  
 $K'(1) = C_1$ ,  $\alpha = K''(1)/K'(1)$   
It is global.

Hence:

$$K(q) = \frac{C_1}{\alpha - 1} (q^\alpha - q); \quad 0 \leq \alpha \leq 2$$

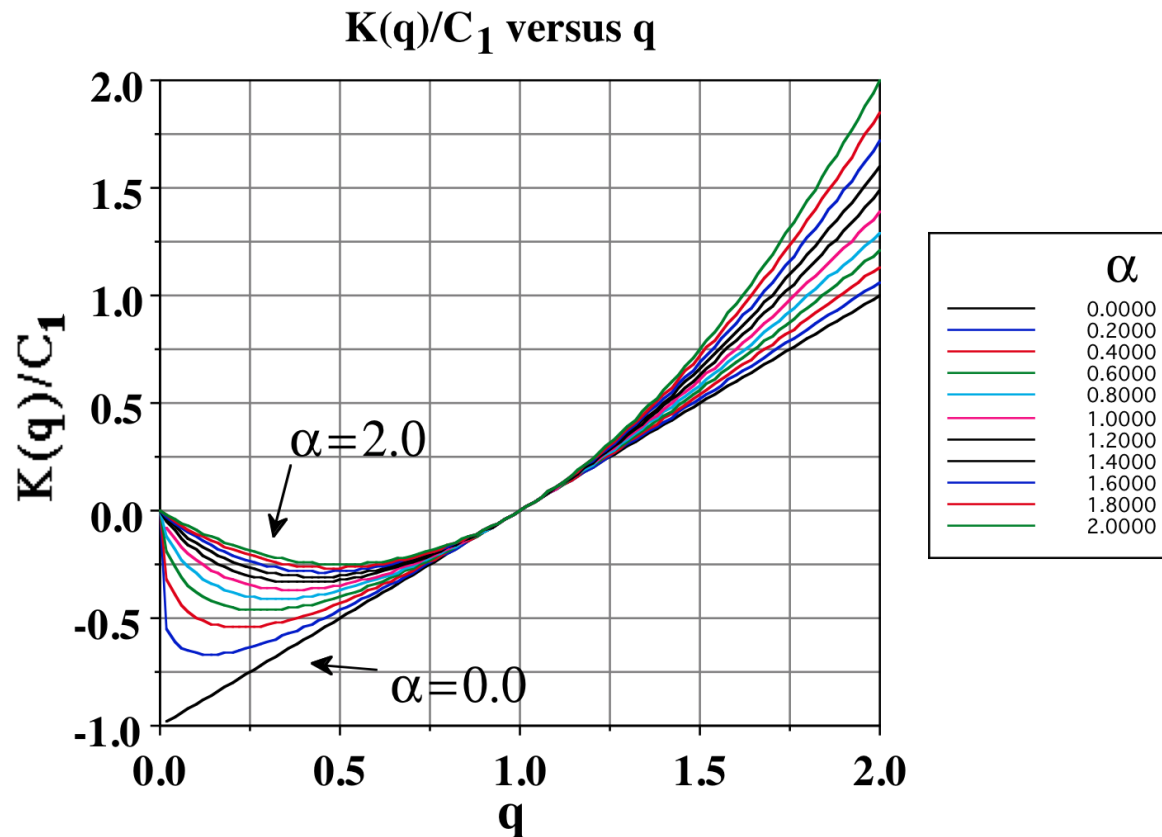
(for  $\alpha = 1$ , using l'Hôpital's rule for the limit  $\alpha \rightarrow 1$ , we have  $C_1 q \log q$ ).

Note that when  $\alpha < 2$ , and  $q < 0$ , then ; this is a consequence of the extreme Lévy tail on the negative (but not positive) fluctuations of  $\log \varepsilon$ . The possibility (even likelihood) of:  $\langle \varepsilon_\lambda^q \rangle \rightarrow \infty$

for  $q < 0$  means that extreme caution should be used when analysing negative moments of empirical data.

# K(q) for universal multifractals

$$K(q)/C_1 = (q^\alpha - q)/(\alpha - 1)$$



Universal  $K(q)/C_1$  as a function of  $q$ , for different  $\alpha$  values from 0 to 2 by increments of  $\Delta\alpha = 0.2$ .

# Data Analysis

# Fluctuation statistics and structure functions

The space-time variability of natural systems, can often be broken up into various “scaling ranges” over which the fluctuations vary in a power law manner with respect to scale. Over these ranges, the fluctuations follow

$$\Delta T = \varphi_{\Delta t} \Delta t^H$$

The flux at resolution  $\Delta t$

Using Fluctuations:

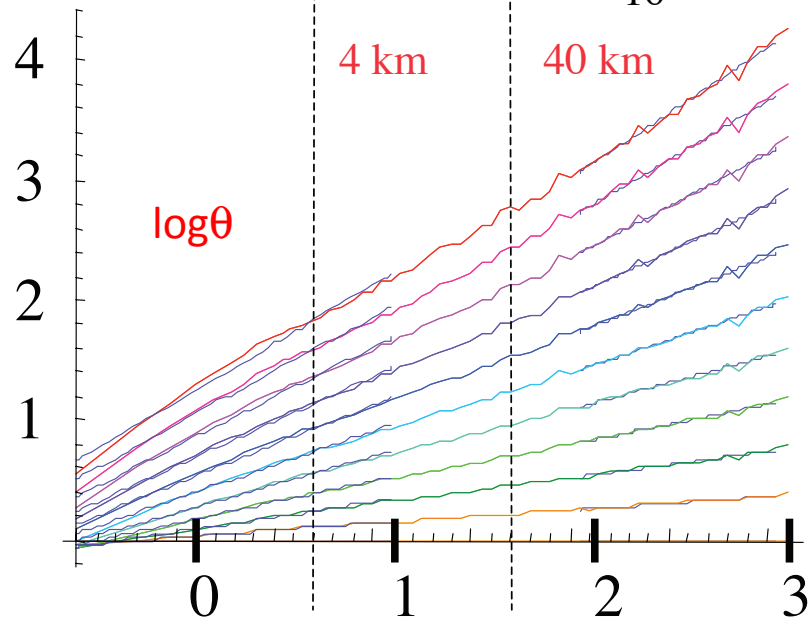
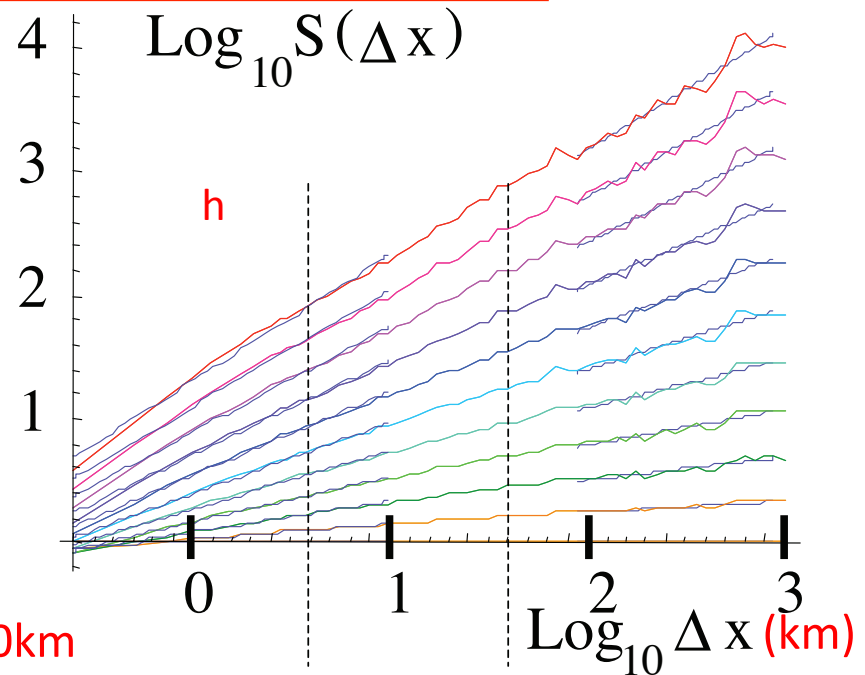
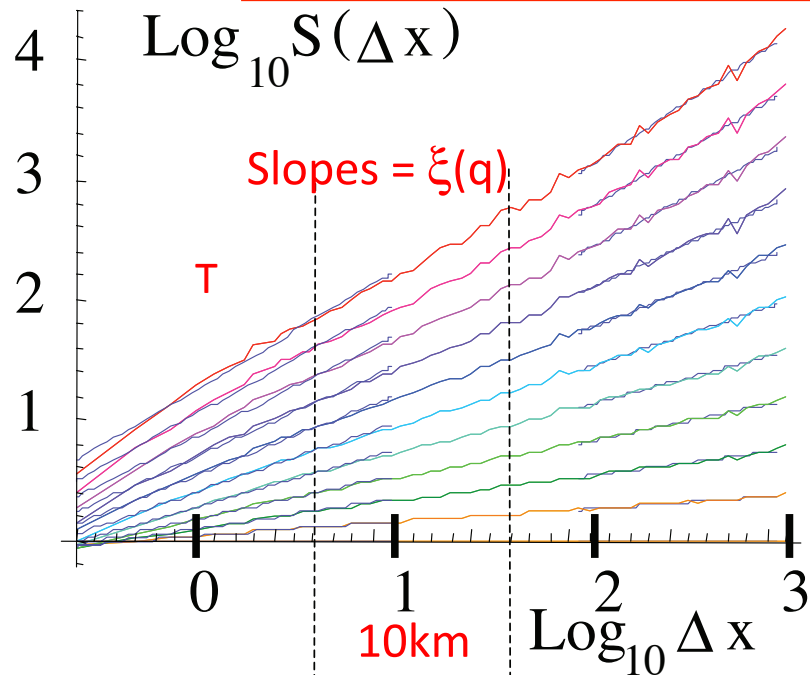
$$S_q(\Delta t) = \langle \Delta T (\Delta t)^q \rangle = \langle \varphi_{\Delta t}^q \rangle \Delta t^{qH} \approx \Delta t^{\xi(q)}; \quad \langle \varphi_{\Delta t}^q \rangle = \left( \frac{\tau}{\Delta t} \right)^{K(q)}; \quad \xi(q) = qH - K(q)$$

(generalized, qth order) Structure function

Hence, we seek H, K(q)

With universality:  $K(q) = \frac{C_1}{\alpha - 1} (q^\alpha - q)$  i.e. we seek H,  $C_1$ ,  $\alpha$

# Aircraft structure function estimates

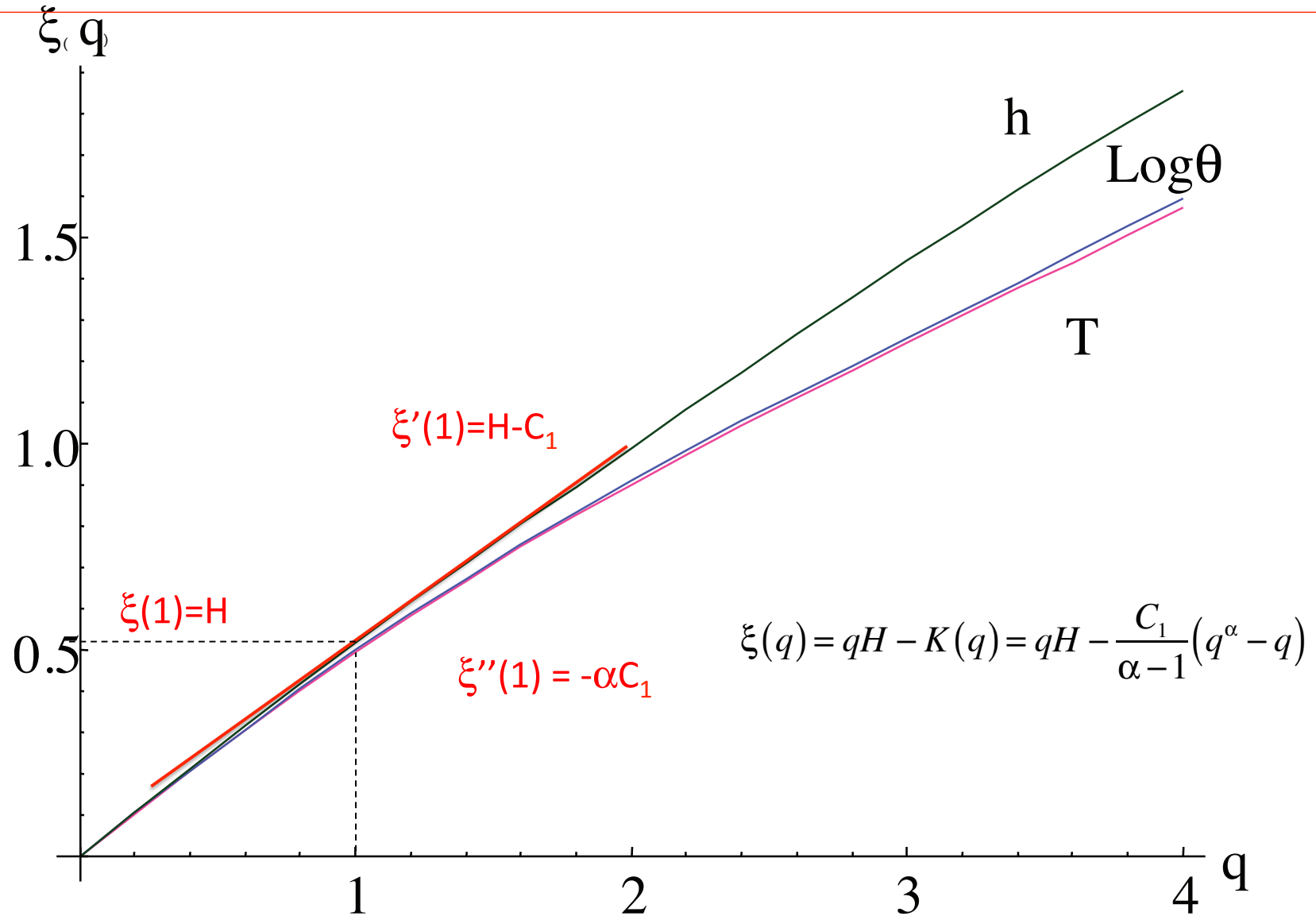


Fluctuations as differences

Temperature (Upper left),  
humidity (upper right), log  
potential temperature (lower  
left)

The structure functions of order  $q = 0.2, 0.4, \dots, 1.8, 2.0$  are shown (from bottom to top). All have been nondimensionalized by dividing by the absolute mean first difference at the finest scale (280 m)

# $\xi(q)$



The structure function exponents for  $T$ ,  $\log\theta$ ,  $h$  from the aircraft data analysed in the previous slide. The exponents were estimated by fitting the structure functions over the “optimal” range 4 – 40 km.

# Difference, Tendency, Haar fluctuations

**Differences:** The difference in temperature between  $t$  and  $t+\Delta t$

**Tendency:** The average of the temperature (with overall mean removed) between  $t$  and  $t+\Delta t$

**Haar:** The difference between the average of the temperature from  $t$  and  $t+\Delta t/2$  and from  $t+\Delta t/2$  and  $t+\Delta t$

**Relations:** When  $1 > H > 0$ : Haar  $\approx$  difference  
When  $0 > H > -1$ : Haar  $\approx$  tendency



# Fluctuations and wavelets

$$\Delta T(\Delta t) = \frac{1}{\Delta t} \int T(t') \Psi\left(\frac{t-t'}{\Delta t}\right) dt'$$

mother wavelet

Difference

$$(\Delta T)_{diff} = T(t + \Delta t / 2) - T(t - \Delta t / 2) \quad \Psi(t) = \delta(t - 1/2) - \delta(t + 1/2)$$

Tendency / Anomaly

$$(\Delta T)_{tend} = \frac{1}{\Delta t} \int_t^{t+\Delta t} T'(t') dt'; \quad T'(t) = T(t) - \overline{T(t)} \quad \Psi(t) = I_{[-1/2, 1/2]}(t) - \frac{I_{[-\tau/2, \tau/2]}(t)}{\tau}; \quad \tau \gg 1 \quad I_{[a,b]}(t) = \begin{cases} 1 & a \leq t \leq b \\ 0 & \text{otherwise} \end{cases}$$

$I$  is the indicator function

Haar

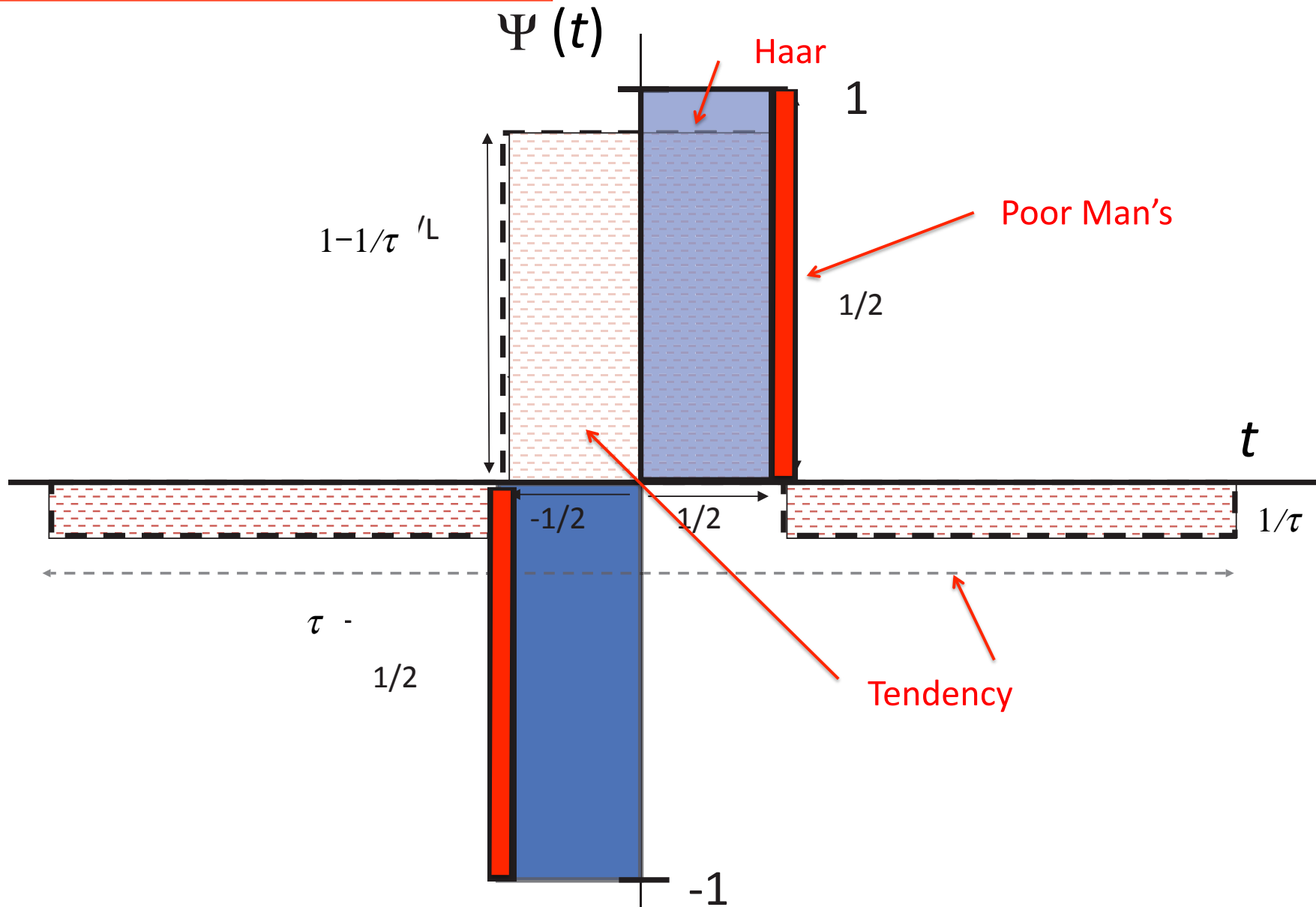
$$(\Delta T)_{Haar} = \frac{2}{\Delta t} \left[ \int_t^{t+\Delta t/2} T(t') dt' - \int_{t+\Delta t/2}^{t+\Delta t} T(t') dt' \right] \quad \Psi(t) = \begin{cases} 1/2; & 0 \leq t < 1/2 \\ -1/2; & -1/2 \leq t < 0 \\ 0; & \text{otherwise} \end{cases}$$

Relation between them:

$$(\Delta T)_{Haar} = \left( \Delta (\Delta T)_{tend} \right)_{diff}$$

# Haar, tendency and poor man's wavelets

$$\Delta T(\Delta t) = \frac{1}{\Delta t} \int T(t') \Psi\left(\frac{t' - t}{\Delta t}\right) dt'$$



# Spectrum of fluctuations

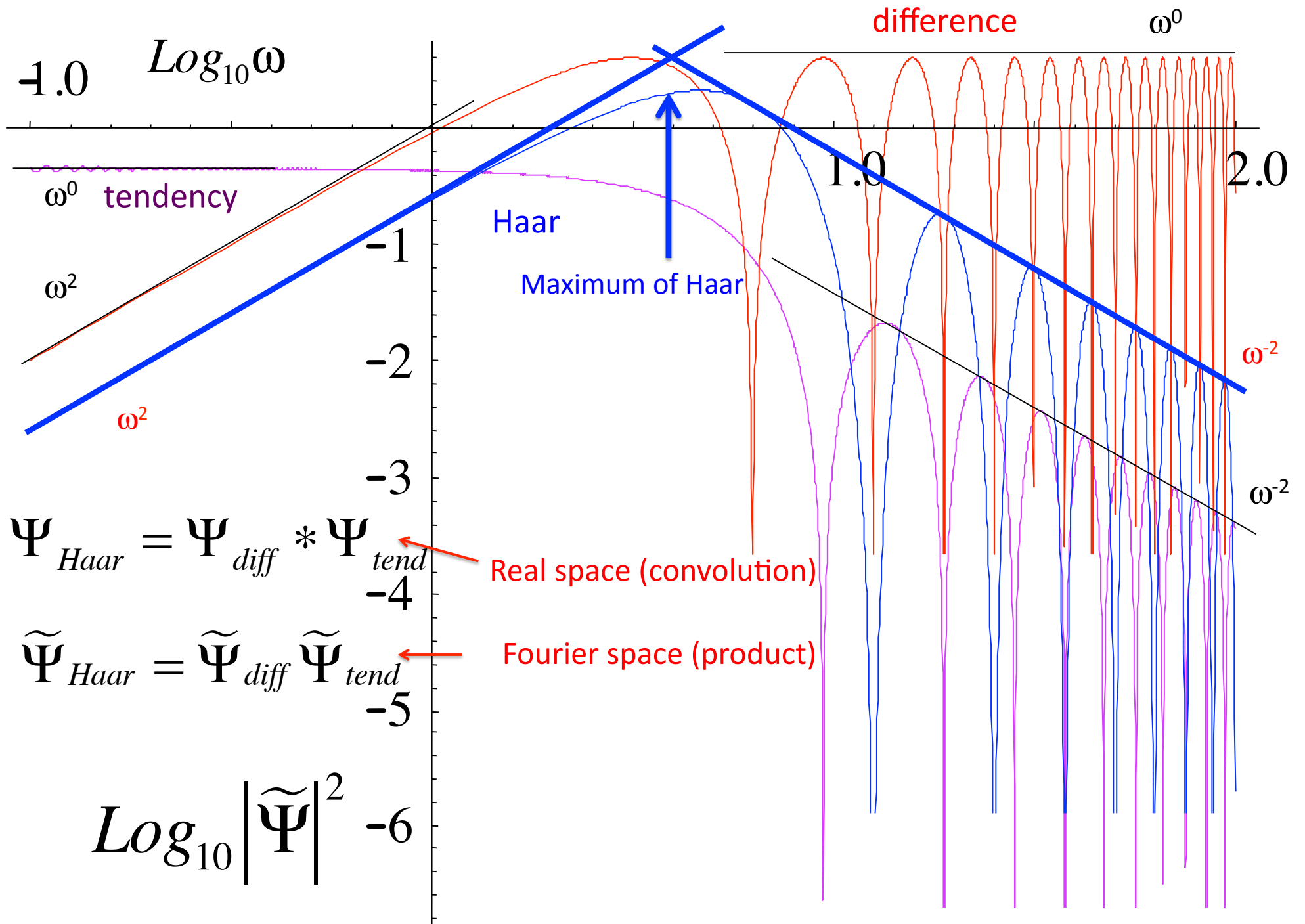
Fluctuations: 
$$\Delta T(\Delta t) = \frac{1}{\Delta t} \int T(t') \Psi\left(\frac{t' - t}{\Delta t}\right) dt'$$

Fourier transforms 
$$\widetilde{\Delta T}(\omega \Delta t) = \widetilde{T}(\omega) \widetilde{\Psi}(\omega)$$

Ensemble averaging  
of modulus  
squared: 
$$\left\langle \left| \widetilde{\Delta T}(\omega \Delta t) \right|^2 \right\rangle = \left\langle \left| \widetilde{T}(\omega) \right|^2 \right\rangle \left| \widetilde{\Psi}(\omega) \right|^2$$

Spectra 
$$E_{\Delta T}(\omega \Delta t) = E_T(\omega) \left| \widetilde{\Psi}(\omega) \right|^2$$

If the maximum of  $\left| \widetilde{\Psi}(\omega) \right|^2$  Occurs at  $\omega_m$ , then the maximum in  $E_{\Delta T}$  may be near  $\omega_m \Delta t$



# Convergence of fluctuation variance

Spectra

$$E_{\Delta T}(\omega \Delta t) = E_T(\omega) \left| \overline{\Psi(\omega)} \right|^2$$

For scaling processes

$$E_T(\omega) \approx \omega^{-(1+2H')}$$
$$\left| \overline{\Psi(\omega)} \right|^2 \approx \begin{cases} \omega^{2H_{low}}; & \omega \rightarrow 0 \\ \omega^{2H_{high}}; & \omega \rightarrow \infty \end{cases}$$

$$\underbrace{\langle \Delta T^2 \rangle = \int_{-\infty}^{\infty} E_{\Delta T}(\omega) d\omega}_{\text{Parseval's theorem}}$$

Converges only if:

$$H_{low} > H' > H_{high}$$

# Various wavelets

Name	Wavelet	Frequency domain	low $\omega$	high $\omega$	H' range
Poor man's (first difference)	$\delta(t-1/2) - \delta(t+1/2)$	$2\sin(\omega/2)$	$\approx \omega$	$\approx 0$	$0 \leq H' \leq 1$
2 <sup>nd</sup> difference	$\frac{1}{2}(\delta(t+1/2) + \delta(t-1/2)) - \delta(t)$	$\sin^2(\omega/4)$	$\approx \omega^2$	$\approx 0$	$0 \leq H' \leq 1$
Tendency	$I_{[-1/2, 1/2]}(t) - \frac{I_{[-\tau/2, \tau/2]}(t)}{\tau}; \quad \tau \gg 1$	$\frac{2}{\omega} \left( \sin\left(\frac{\omega}{2}\right) - \tau^{-1} \sin\left(\frac{\omega\tau}{2}\right) \right)$	$\frac{2\sin(\frac{\omega\tau}{2})}{\omega\tau} \approx 0; \quad \omega\tau \gg 1$	$\approx \omega^{-1}$	$-1 \leq H' \leq 0$
Haar	$\psi(t) = \begin{cases} 1/2; & 0 \leq t < 1/2 \\ -1/2; & -1/2 \leq t < 0 \\ 0; & \text{otherwise} \end{cases}$	$2i\omega^{-1} \sin^2\left(\frac{\omega}{4}\right)$	$\approx \omega$	$\approx \omega^{-1}$	$-1 \leq H' \leq 1$
Quadratic Haar	$\psi(t) = \begin{cases} -1/3 & 1/3 < t < 1 \\ 2/3; & -1/3 \leq t \leq 1/3 \\ -1/3; & -1 \leq t < -1/3 \\ 0; & \text{otherwise} \end{cases}$	$\frac{2}{3\omega} \left( 3\sin\frac{\omega}{3} - \sin\omega \right)$	$\approx \omega^2$	$\approx \omega^{-1}$	$-1 \leq H' \leq 2$
First derivative Gaussian	$\Psi(t) \propto \frac{d}{dt} e^{-t^2/2}$	$\omega e^{-\omega^2/2}$	$\approx \omega$	$e^{-\omega^2/2}$	$-\infty \leq H' \leq 1$
Mexican Hat	$\Psi(t) \propto \frac{d^2}{dt^2} e^{-t^2/2}$	$\omega^2 e^{-\omega^2/2}$	$\approx \omega^2$	$e^{-\omega^2/2}$	$-\infty \leq H' \leq 2$

Range of exponents over which average fluctuations at scale  $\Delta t$  corresponds to frequency  $1/\Delta t$

Fluctuation  $\langle \Delta I \rangle = \langle \varphi \rangle \Delta t^H = \text{constant}$

$E(\omega) = \langle |\tilde{I}(\omega)|^2 \rangle = \omega^{-\beta}$

$\beta = 1 + 2H - K(2)$

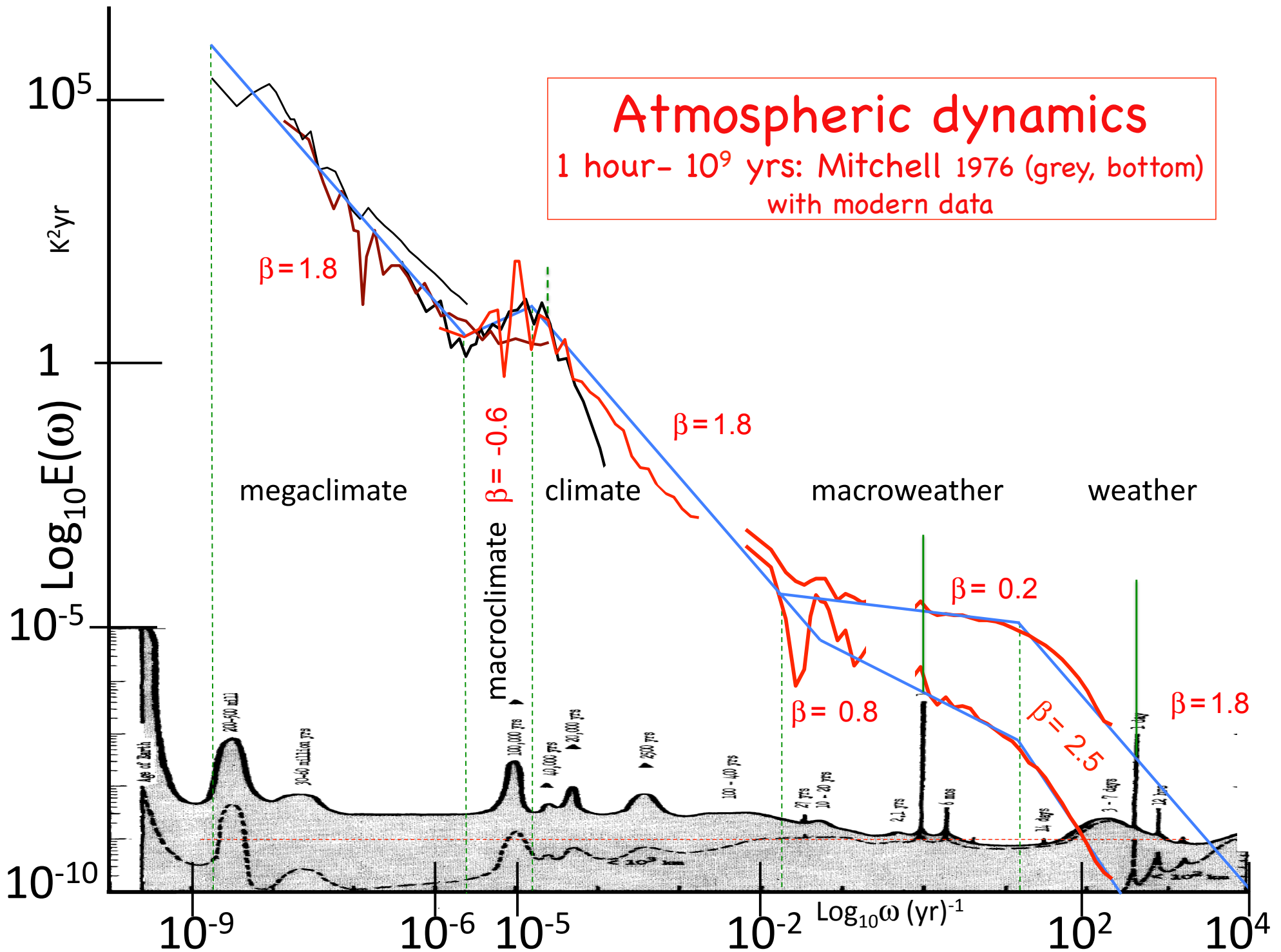
Statistic	Range of H	Range of $\beta$	Comment
Spectrum	$-\infty < H < \infty$	$-\infty < \beta < \infty$	$E(\omega) \approx \omega^{-\beta}$
Difference	$0 < H < 1$	$1 < \beta + K(2) < 3$	"Poor man's wavelet"
Tendency Fluctuation	$-1 < H < 0$	$-1 < \beta + K(2) < 1$	Average with overall mean removed (standard deviation= "Climactogram", also called the "Aggregated Standard Deviation")
Haar	$-1 < H < 1$	$-1 < \beta + K(2) < 3$	Difference of means of first and second halves of interval
Detrended Fluctuation Analysis (DFA, polynomial order n)	$-1 < H < (n+1)$	$-1 < \beta + K(2) < 3+2n$	Also multifractal extension (MFDFA), usually linear: n=1, <b>Not a wavelet</b>
Mexican Hat Wavelet	$-\infty < H < 2$	$-\infty < \beta + K(2) < 5$	2 <sup>nd</sup> Derivative of a Gaussian
Generalized Haar	$-m < H < n$	$1-2m < \beta + K(2) < 3+2n$	Interpretation not simple

Multifractal "correction"  
 $H' = H - K(2)/2$

Simple interpretation

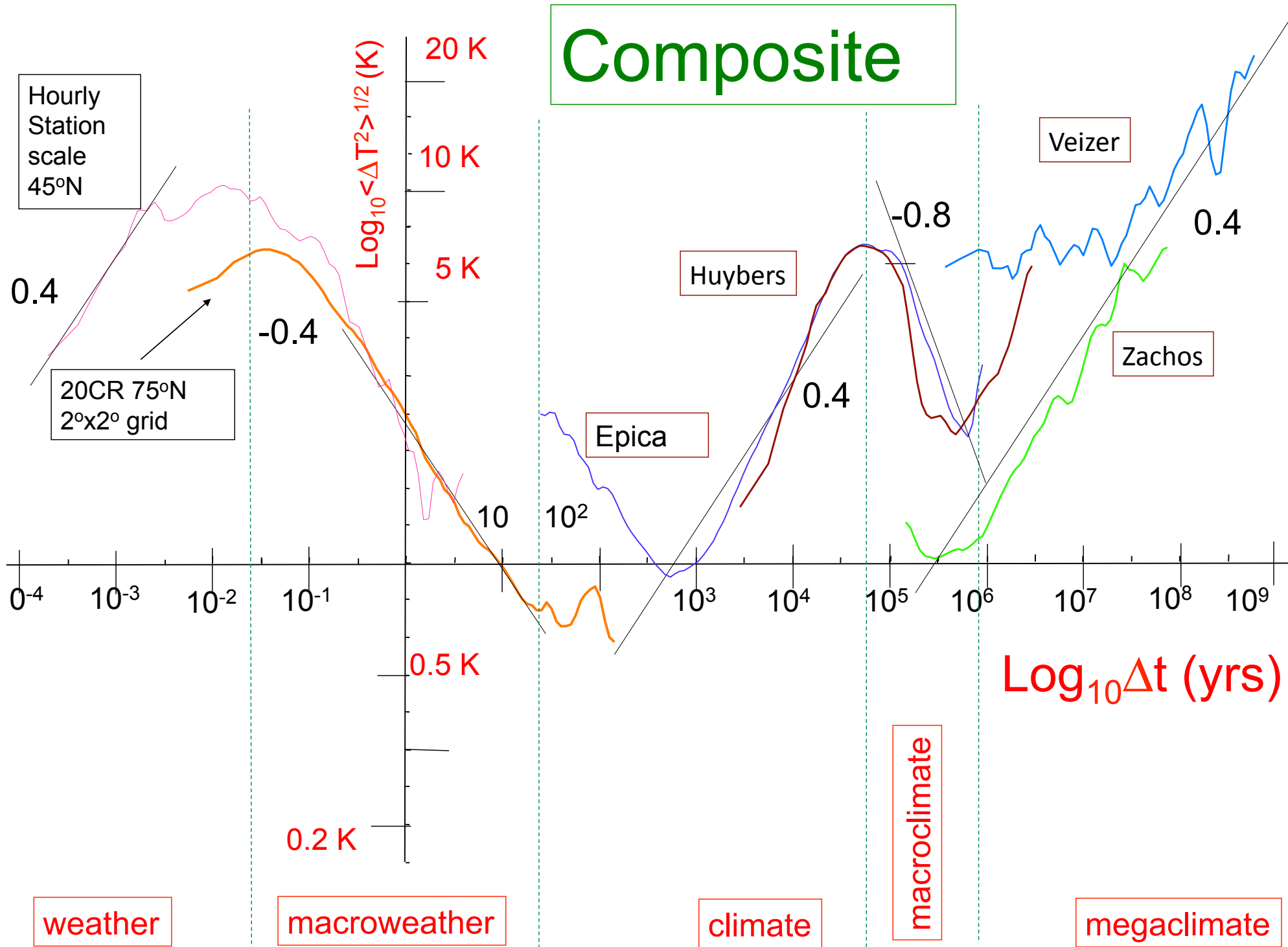
# Atmospheric dynamics

1 hour -  $10^9$  yrs: Mitchell 1976 (grey, bottom)  
with modern data





# Composite



$$\langle \Delta T (\Delta t) \rangle \propto \Delta t^H$$

$H \approx 0.4$

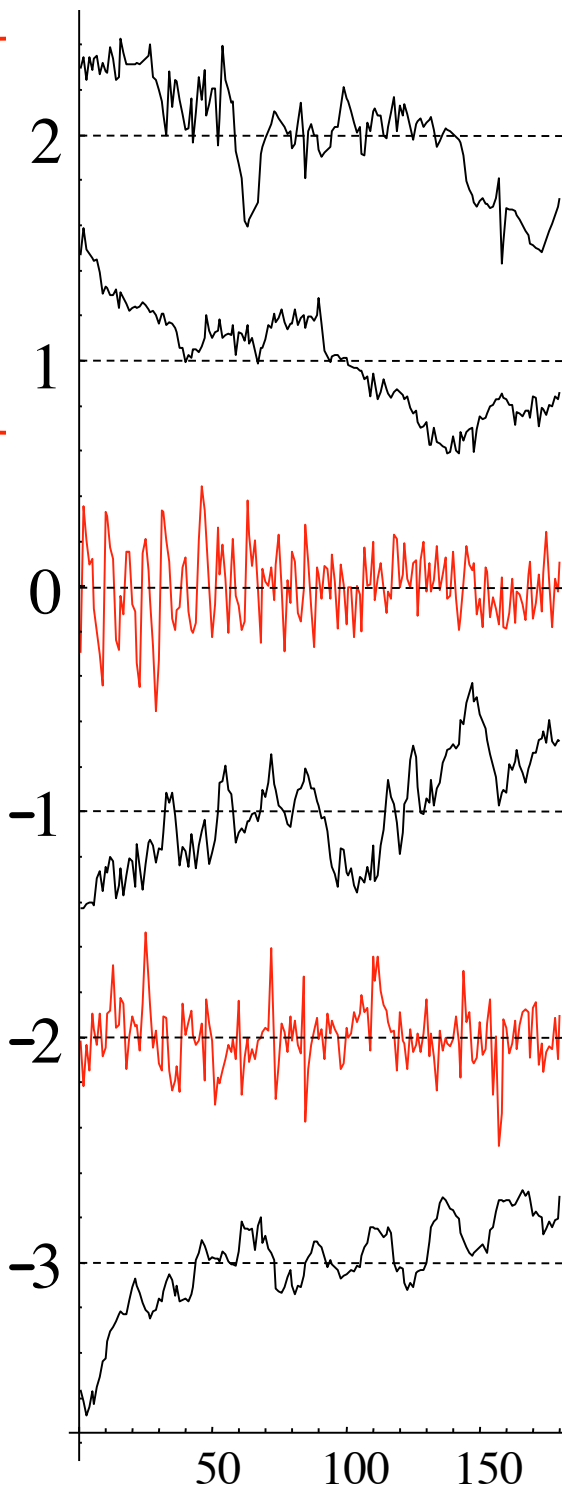
$H \approx -0.8$

$H \approx 0.4$

$H \approx -0.4$

$H \approx 0.4$

$T/\Delta T_{\max}$



**Megaclimate**

Veizer: 290 Mys - 511 Myrs BP (1.23 Myr)

**Megaclimate**

Zachos: 0-67 Myrs (370 kyr)

**Macroclimate**

Huybers: 0-2.56 Myrs (14 kyrs)

**Climate**

Epica: 25-97 BP kyrs (400 yrs)

**Macroweather**

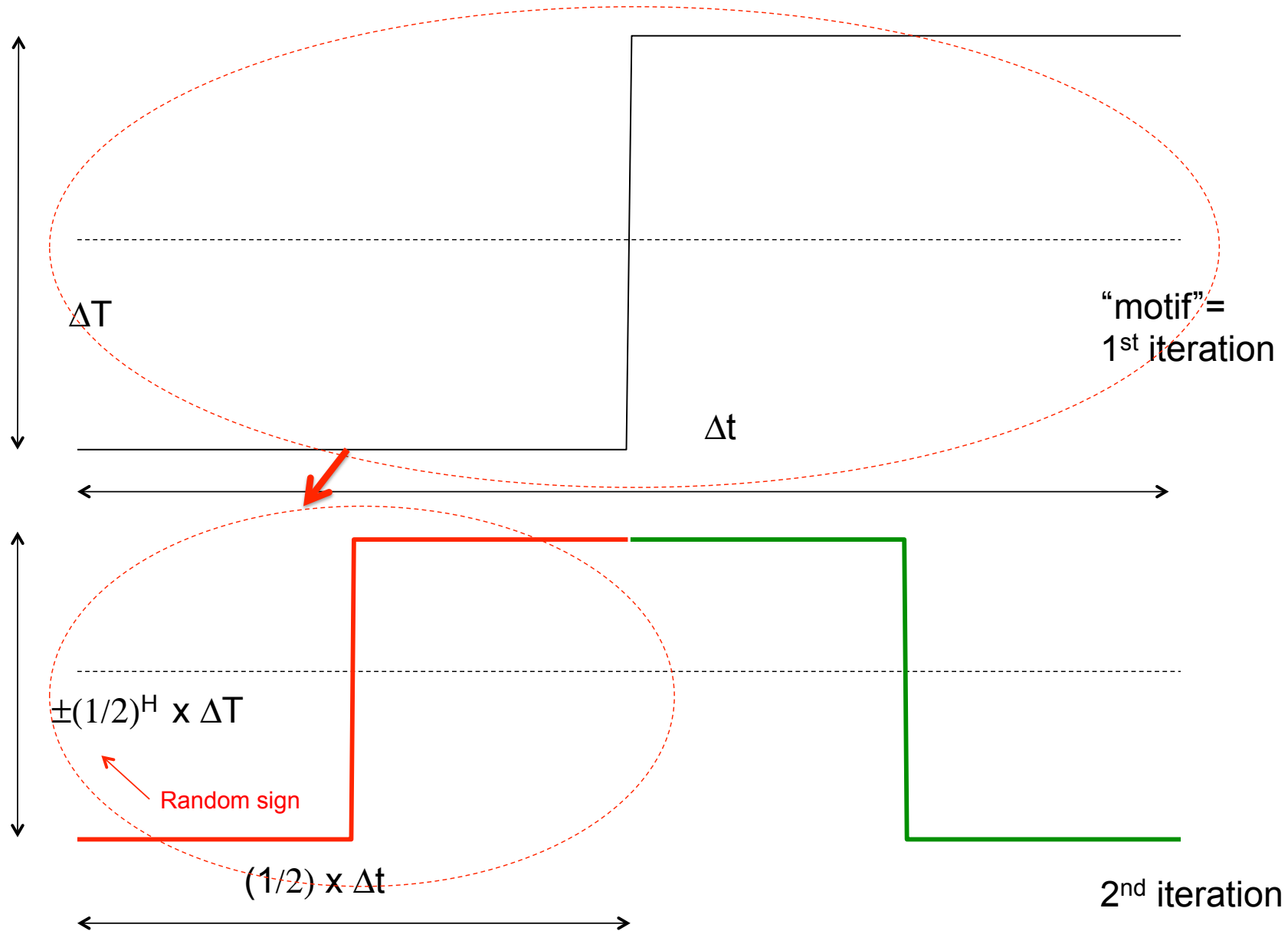
Berkeley: 1880-1895 AD (1 month)

**Weather**

Lander Wy.: July 4-July 11, 2005 (1 hour)

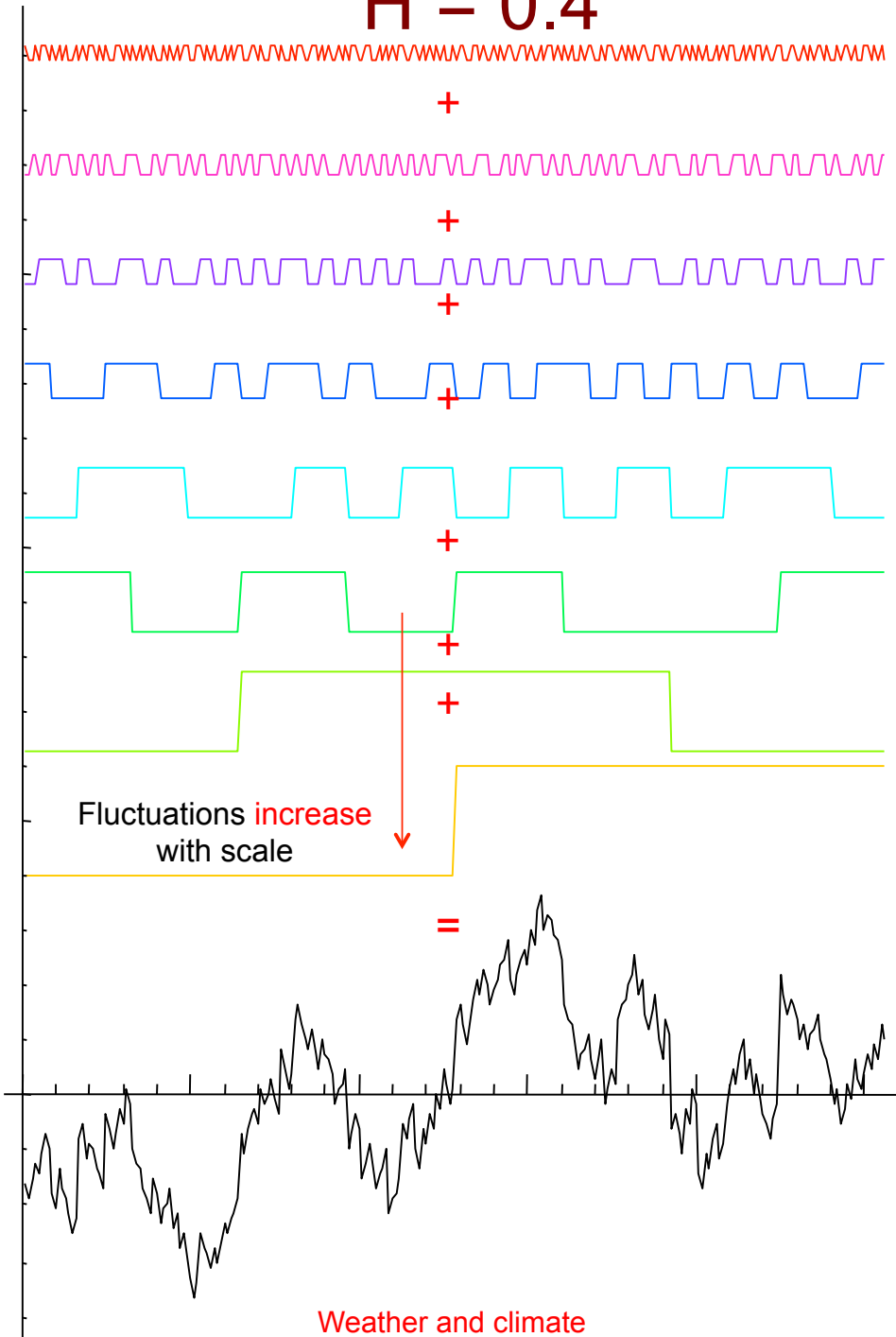
# Understanding the Fluctuation exponent $H$

# The fractal H model



$H = 0.4$

$H = -0.4$



Scale increasing

