Scale, scaling and multifractals in geophysics

Part 3:

Data Analysis, Codimensions

Course at U. Paris Sud, May 6, 7 2014

7 May, 2

Spectral analysis

 $R(\Delta t) = \langle T(t)T(t - \Delta t) \rangle$ autocorrelation $S(\Delta t) = \left\langle \left(T(t) - T(t - \Delta t) \right)^2 \right\rangle = \left\langle \Delta T(\Delta t)^2 \right\rangle$ Classical (q=2) structure function $S(\Delta t) = 2(R(0) - R(\Delta t))$ Relation between them $\left\langle \widetilde{T}(\omega)\widetilde{T}^{*}(\omega')\right\rangle = \delta(\omega+\omega')E(\omega); \quad \widetilde{T(\omega)} = \int_{-\infty}^{\infty} T(t)e^{-i\omega t} dt$ Spectral analysis (spectrum $E(\omega)$) Statistics independent of time (stationarity) Note: for real T(t), we have: $\widetilde{T}(\omega) = \widetilde{T}^*(-\omega)$ hence $E(\omega) \propto \langle |\widetilde{T(\omega)}|^2 \rangle$ $R(\Delta t) = \int E(\omega) e^{i\omega\Delta t} d\omega$ Wiener-Khintchine theorem $R(0) = \int_{0}^{\infty} \tilde{E}(\omega) d\omega$ Real space- $S(\Delta t) = 2 \int_{-\infty}^{\infty} (1 - e^{i\omega\Delta t}) E(\omega) d\omega$ Real space-Fourier space relation Course at U. Paris Sud, May 6, 7 2014

Tauberian theorem

If the spectrum is of power $E(\omega) \approx \omega^{-\beta}$ law form:

Then:

—∞

$$E(\omega / \lambda) = \lambda^{\beta} E(\omega)$$

Using:

 ∞

 $\omega = \omega' / \lambda$

$$S(\Delta t) = 2 \int_{-\infty}^{\infty} (1 - e^{i\omega'\Delta t/\lambda}) \lambda^{-\beta} E(\omega'/\lambda) \frac{d\omega'}{\lambda}$$
$$S(\lambda \Delta t) = 2 \int_{-\infty}^{\infty} (1 - e^{i\omega'\Delta t}) E(\omega') \lambda^{-1+\beta} d\omega' = \lambda^{-1+\beta} S(\Delta t)$$

Hence:

$$S(\Delta t) \approx \Delta t^{2H'}; \quad H' = (\beta - 1)/2$$

We conclude:

POWER LAWS <->F.T. POWER LAWS

For convergence of the integral: $1 < \beta < 3$ ($0 \le H' \le 1$) for $S(\Delta t)$, $1 > \beta > -1$; for $R(\Delta t)$ (-1/2 < H' < 0)

Empirical analysis: Estimating fluxes from the fluctuations

Multifractal cascade equation:

$$\left\langle \varphi_{\lambda}^{q} \right\rangle = \lambda^{K(q)}$$

Fluctuations:

Estimating the fluxes from the fluctuations

$$\Delta T = \varphi_{\Delta t} \Delta t^{H}$$

$$\varphi_{\lambda}' = \frac{\varphi_{\lambda}}{\langle \varphi_{\lambda} \rangle} \approx \frac{\Delta T (\Delta t)}{\langle \Delta T (\Delta t) \rangle}; \quad \lambda = \frac{\tau}{\Delta t}$$
outer cascade scale

Normalized flux at resolution λ

 $M_q = \left\langle \varphi_{\lambda}^{\prime q} \right\rangle \leftarrow$ \uparrow "Trace moments"

Estimate at finest resolution, then degrade to intermediate resolutions by averaging



TRMM satellite data, ≈1000 orbits

Energy budget







Vertical cascades: lidar backscatter



From 10 airborne lidar cross-sections near Vancouver B.C.



Vertical cascade









Horizontal spatial Scaling exponents

		C1	α	н	β	L _{eff}
State variables	u, v w T h z	0.09 (0.12) 0.11, (0.08) 0.09 (0.09)	1.9 (1.9) 1.8 1.8 (1.9)	1/3, (0.77) (—0.14) 0.50, (0.77) 0.51 (1.26)	1.6, (2.4) (0.4) 1.9, (2.4) 1.9 (3.3)	(14 000) (15 000) 5000 (19 000) 10 000 (60 000)
Precipitation	R	0.4	1.5	0.00	0.2	32 000
Passive scalars	Aerosol concentration	0.08	1.8	0.33	1.6	25 000
Radiances	Infrared Visible Passive microwave	0.08 0.08 0.1–0.26	1.5 1.5 1.5	0.3 0.2 0.25–0.5	1.5 1.5 1.3–1.6	15 000 10 000 5000–15 000
Topography	Altitude	0.12	1.8	0.7	2.1	20 000
Sea surface temperature	SST (see Table 8.2)	0.12	1.9	0.50	1.8	16 000

$$\Delta I = \varphi \Delta x^{H} \quad \left\langle \varphi_{\lambda}^{q} \right\rangle = \lambda^{K(q)} \quad \lambda = L_{eff} / \Delta x \quad K(q) = \frac{C_{1}}{\alpha - 1} (q^{\alpha} - q) \quad E(k) \approx k^{-\beta}$$

Surface, solid earth exponents

	C ₁	α	Н	β
Rock Density (vertical)	0.045	2.0	0.08	1.07
Magnetic susceptibility (vertical)	0.11	2.0	0.17	1.12
Topography	0.12	1.8	0.7	2.1
Vegetation index	0.064	2.0	0.16	1.19
Soil moisture index	0.053	2.0	0.14	1.17

$$\Delta I = \varphi \Delta x^{H} \quad \left\langle \varphi_{\lambda}^{q} \right\rangle = \lambda^{K(q)} \quad \lambda = L_{eff} / \Delta x \quad K(q) = \frac{C_{1}}{\alpha - 1} \left(q^{\alpha} - q \right) \quad E(k) \approx k^{-\beta}$$

Probabilities and codimensions

Revisiting the α Model



Revisiting the α Model

The α model is a binomial process:

$$Pr(\mu \varepsilon = \lambda_0^{\gamma_+}) = \lambda_0^{-c} \qquad (>1 \implies INCREASE)$$
$$Pr(\mu \varepsilon = \lambda_0^{\gamma_-}) = 1 - \lambda_0^{-c} \qquad (<1 \implies DECREASE)$$

where γ_+ , γ_- correspond to boosts and decreases respectively, the β model being the special case where $\gamma_- = -\infty$ and $\gamma_+ = c$ (due to conservation $\langle \mu \epsilon \rangle = 1$, there are only two free parameters): $\lambda_0^{\gamma_+ - c} + \lambda_0^{\gamma_-} (1 - \lambda_0^{-c}) = 1$ Conservation constraint

Taking $\gamma_- > -\infty$, the pure orders of singularity γ_- and γ_+ lead to the appearance of mixed orders of singularity, of different orders $\gamma (\gamma_- \le \gamma \le \gamma_+)$.

α Model after 2 steps

What is the behaviour as the number of cascade steps, $n \rightarrow \infty$? Consider two steps of the process, the various probabilities and random factors are:

Two steps: an
equivalent 3
state model
with $\lambda = \lambda_0^2$ $\Pr(\mu \epsilon = \lambda_0^{2\gamma_+}) = \lambda_0^{-2c}$
 $\Pr(\mu \epsilon = \lambda_0^{\gamma_+ + \gamma_-}) = 2\lambda_0^{-c}(1 - \lambda_0^{-c})$ (two boosts) $\Pr(\mu \epsilon = \lambda_0^{2\gamma_-}) = (1 - \lambda_0^{-c})^2$ (two decrease)

Rewriting:
$$\Pr(\mu \varepsilon = (\lambda_0^2)^{\gamma_+}) = (\lambda_0^2)^{-c} \qquad \text{(one large)}$$
$$\Pr(\mu \varepsilon = (\lambda_0^2)^{(\gamma_+ + \gamma_-)/2}) = 2(\lambda_0^2)^{-c/2} - 2(\lambda_0^2)^{-c} \qquad \text{(intermediate)}$$
$$\Pr(\mu \varepsilon = (\lambda_0^2)^{\gamma_-}) = 1 - 2(\lambda_0^2)^{-c/2} + (\lambda_0^2)^{-c} \qquad \text{(large decrease)}$$

α Model after n steps

Iterating this procedure, after $n = n_+ + n_-$ steps we find:

$$\gamma_{n_{+},n_{-}} = \frac{n_{+}\gamma_{+} + n_{-}\gamma_{-}}{n_{+} + n_{-}}, \qquad n_{+} = 1,...,n$$
$$\Pr\left(\mu\varepsilon = \left(\lambda_{0}^{n}\right)^{\gamma_{n_{+},n_{-}}}\right) = \binom{n}{n_{+}} \left(\lambda_{0}^{n}\right)^{-cn_{+}/n} \left(1 - \left(\lambda_{0}^{n}\right)^{-c/n}\right)^{n_{+}}$$

where $\binom{n}{n_+}$ is the number of combinations of *n* objects taken n^+ at a time. This implies that we may write: $\Pr(\mathbf{s} \ge (\lambda^n)^{\gamma_i}) - \sum p_-(\lambda^n)^{-c_{i,j}}$

$$\Pr\left(\varepsilon_{\lambda_0^n} \ge \left(\lambda_0^n\right)^{\gamma_i}\right) = \sum_j p_{i,j} \left(\lambda_0^n\right)^{-c_{i,j}}$$

The p_{ij} 's are the "submultiplicities" (the prefactors in the above), c_{ij} are the corresponding exponents ("subcodimensions") and λ_0^n is the total ratio of scales from the outer scale to the smallest scale. Notice that the requirement that $\langle \mu \varepsilon \rangle = 1$ implies that some of the λ^{γ_i} are >1.

Values and singularities



A schematic illustration of a multifractal field analyzed over a scale ratio λ , with two scaling thresholds λ^{γ_1} and λ^{γ_2} ., corresponding to two orders of singularity : $\gamma_2 > \gamma_1$.



The Codimension Multifractal Formalism

Codimension of Singularities $c(\gamma)$ and its relation to K(q)

We now derive the basic connection between $c(\gamma)$ and the moment scaling exponent K(q). To relate the two; write the expression for the moments in terms of the probability density of the singularities:

$$p(\gamma) = \left| \frac{d \operatorname{Pr}}{d \gamma} \right| \sim c'(\gamma) (\log \lambda) \lambda^{-c(\gamma)} \sim \lambda^{-c(\gamma)}$$

Relation probability density and distribution: $Pr(\gamma' > \gamma) = \int_{\alpha}^{\beta} p(\gamma') d\gamma'$

(where we have absorbed the $c'(\gamma)\log\lambda$ factor into the "~" symbol since it is slowly varying, subexponential). This yields:

$$\langle \varepsilon_{\lambda}^{q} \rangle = \int d\Pr(\varepsilon_{\lambda}) \varepsilon_{\lambda}^{q} \approx \int d\gamma \lambda^{-c(\gamma)} \lambda^{q\gamma} \qquad \Pr(\varepsilon_{\lambda} > \lambda^{\gamma}) = \int_{\lambda^{\gamma}}^{\infty} p(\varepsilon_{\lambda}) d\varepsilon_{\lambda}$$

where we have used $\varepsilon_{\lambda} = \lambda^{\gamma}$ (this is just a change of variables ε_{λ} for γ , λ is a fixed parameter). Hence:

$$\langle \varepsilon_{\lambda}^{q} \rangle = \lambda^{K(q)} = e^{K(q)\log\lambda} \approx \int_{-\infty}^{\infty} d\gamma e^{\xi f(\gamma)}; \quad \xi = \log\lambda; \quad f(\gamma) = q\gamma - c(\gamma); \quad \lambda >> 1$$

Legendre transform

We need an asymptotic expansion of an integral with integrand of the form:

 $\exp(\xi f(\gamma)); \quad \xi = \log \lambda \text{ is a large parameter}, \qquad f(\gamma) = q\gamma - c(\gamma).$

"Steepest descents" method shows that for $\xi = \log \lambda >> 1$, the integral is dominated by the γ which yields the maximum:

$$\int_{-\infty}^{\infty} e^{\xi f(\gamma)} d\gamma \approx e^{\xi \max_{\gamma} (f(\gamma))}$$

so that:

$$\langle \varepsilon_{\lambda}^{q} \rangle = e^{\xi K(q)} \approx e^{\xi \max_{\gamma} (q\gamma - c(\gamma))}; \quad \xi = \log \lambda$$

hence:

$$K(q) = \max_{\gamma} (q\gamma - c(\gamma))$$
 Legendre transform

This relation between K(q) and $c(\gamma)$ is called a "Legendre transform" (Parisi and Frisch, 1985).

Inverse Legendre transform: $c(\gamma)$

We can also invert the relation to obtain $c(\gamma)$ from K(q): aLegendre transform is equal to its inverse, hence we conclude:

The γ which for a given q maximizes $q\gamma - c(\gamma)$ is γ_q and is the solution of $c'(\gamma_q) = q$. Similarly, the value of q which for given γ maximizes $q\gamma - K(q)$ is q_{γ} so that:



 $c(\gamma)$ versus showing the tangent line $c'(\gamma_q) = q$ with the corresponding chord. Note that the equation is the same as $\gamma_q = K'(q)$.

Graphical Legendre transform



K(q) versus q showing the tangent line $K'(q_{\gamma}) = \gamma$ with the corresponding chord.

Properties of codimension functions

 $c(\gamma)$ is the statistical scaling exponent characterizing how its probability changes with scale.

1) The first obvious property is that due to its very definition $c(\gamma)$ is an increasing function of γ : $c'(\gamma) > 0$.

Reason: increasing γ must decrease the probability:

$$\Pr(\varepsilon_{\lambda} > \lambda^{\gamma}) \approx \lambda^{-c(\gamma)}$$

1) Another fundamental property which follows directly from the Legendre relation with K(q), is that $c(\gamma)$ must be convex: $c''(\gamma) > 0$.

Reason: K(q) must be convex and the Legendre transform of a convex function is convex



- 1) First, applying $K'(q) = \gamma$ we find $K'(1) = \gamma_1$ where is the singularity giving the dominant contribution to the mean (the q = 1 moment). We have already defined $C_1 = K'(1)$, so that this implies $C_1 = \gamma_1$; the Legendre relation thus justifies the name "codimension of the mean" for C_1 .
- 2) Also at q = 1 we have K(1) = 0 (due to the scale by scale conservation of the flux) so that $C_1 = c(C_1)$ (this is a fixed point relation). C_1 is thus simultaneously the codimension of the mean of the process and the order of singularity giving the dominant contribution to the mean.

Reason:
$$K(1) = \max_{\gamma} (\gamma - c(\gamma))$$
 Now, note that the γ that maximizes this is $\gamma_1 = C_1$ and $K(1) = 0$

1) Finally, applying $c'(\gamma) = q$ we obtain $c'(C_1) = 1$ so that the curve $c(\gamma)$ is also tangent to the line x = y (the bisectrix). If the process is observed on a space of dimension *d*, it must satisfy $d \ge C_1$, otherwise, following the above, the mean will be so sparse that the process will (almost surely) be zero everywhere; it will be "degenerate". We will see that when $C_1 > d$ that the ensemble mean of the spatial averages (the dressed mean) cannot converge. **Course at U. Paris Sud**, May 6, 7 2014 Codimensions of Universal multifractals, cascades

$$K(q) = \frac{C_1}{\alpha - 1} (q^{\alpha} - q); \qquad \alpha \neq 1 \qquad \text{Universal multifractal K(q)}$$

$$K(q) = C_1 q Log(q); \quad \alpha = 1$$

Valid for $0 \le \alpha \le 2$; however, *K* diverges for all q < 0 except in the special ("log-normal") case $\alpha = 2$. To obtain the corresponding $c(\gamma)$, one can simply take the Legendre transformation to obtain Add: to obtain the $\alpha = 1$ case, just take limit as $\alpha -> 1$.

$$c(\gamma) = C_1 \left(\frac{\gamma}{C_1 \alpha'} + \frac{1}{\alpha}\right)^{\alpha'}; \quad \alpha \neq 1; \quad 1/\alpha' + 1/\alpha = 1$$
$$c(\gamma) = C_1 e^{\left(\frac{\gamma}{C_1} - 1\right)}; \quad \alpha = 1$$
Universal multifractal c(\gamma)

Universal $c(\gamma)$

 $c(\gamma)/C$ versus γ/C_1 2.0 α 1.5 0.0000 0.2000 0.4000 $\alpha = 0$ 0.6000 0.8000 $c(\gamma)/C_1$ 1.0000 1.2000 1.4000 1.0 1.6000 1.8000 2.0000 0.5 $\alpha = 2.0$ 0.0 -0.5 0.5 -1.0 0.0 1.0 1.5 2.0 γ/C_1

Universal $c(\gamma)$ vs γ , for different $\alpha = 0$ to 2 by increment $\Delta \alpha = 0.2$.

Note that since α ' changes sign at $\alpha = 1$, for $\alpha < 1$, there is a maximum order of singularity $\gamma_{max} = C_1/(1-\alpha)$ so that the cascade singularities are "bounded", whereas for $\alpha > 1$, there is on the contrary a minimum order $\gamma_{min} = -C_1/(\alpha-1)$ below which the prefactors dominate ($c(\gamma) = 0$ for $\gamma < \gamma_{min}$) but the singularities are unbounded.

α <1, α >1 cases: bounded, unbounded singularities







This shows 11 independent realizations of $\alpha = 0.2$, $C_1 = 0.1$ indicating the huge realization to realization variability : the bottom realization is not an outlier! no to so impressive with the only exception of a big spike !



Ten independent realizations of $\alpha = 1.9$, $C_1 = 0.1$, again notice the large realization to realization variability.

This shows isotropic realizations in two dimensions with $\alpha = 0.4, 1.2, 2$, (top to bottom) and $C_1 = 0.05, 0.15$ (left to right). The random seed is the same so as to make clear the change in structures as the parameters are changed. The low α simulations are dominated by frequent very low values; the "Lévy holes". The vertical scales are not the same. misleading, we need to find something else..

It's too late to change the name... and if so, to what?



Direct empirical estimation of $c(\gamma)$: the probability distribution multiple scaling (PDMS) technique

$$\Pr(\varepsilon_{\lambda} > \lambda^{\gamma}) \approx \lambda^{-c(\gamma)}$$

Start from the fundamental defining equation, take logs of both sides and rewrite it as follows:

$$Log(Pr(\varepsilon_{\lambda} > \lambda^{\gamma}) = -c(\gamma)Log(\lambda) + o(1 / Log(\lambda))O(\gamma)$$

corresponds to the logarithm of slowly varying factors that are hidden in the " \approx " sign.

The singularity is estimated from the fluxes by:

$$\gamma = \frac{\log(\varepsilon_{\lambda})}{\log \lambda}$$

Probability Distribution Multiple Scaling technique

Aircraft at 200mb: 24 flight legs, each 4000 points long, 280 *m* resolution (i.e. 1120 *km*), dynamic variables



 $c(\gamma)$ estimated from the PDMS method $c(\gamma) \approx -\log Pr/\log \lambda$ are shown for resolution degraded by factors of 2 from 280 *m* to \approx 36 *km* (longest to shortest curves). For reference, lines of slope 3 (top row) and 5 (bottom row) are given corresponding to power law probability distributions with the given exponents.

Thermodynamic variables



The reference lines all have slopes of 5 Course at U. Paris Sud, May 6, 7 2014

